

# Party Polarization, Ideological Sorting and the Emergence of the U.S. Partisan Gender Gap\*

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## Abstract

We argue that the modern American partisan gender gap—the tendency of men to identify more as Republicans and less as Democrats than women—emerged largely because of mass-level ideological party sorting. As the two major U.S. political parties ideologically polarized at the elite level, the public gradually perceived this polarization and better sorted themselves into the parties that matched their policy preferences. Stable preexisting policy differences between men and women caused this sorting to generate the modern U.S. partisan gender gap. Because education is positively associated with awareness of elite party polarization, the partisan gender gap developed earlier and is consistently larger among those with college degrees. We find support for this argument from decades of American National Election Studies data and a new large dataset of decades of pooled individual-level Gallup survey responses.

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# 1 Introduction

It is well known that, in the United States, men now identify more as Republicans and less as Democrats than women do (e.g., Mansbridge, 1986; Mueller, 1988; Anderson, 1997; Wolbrecht, 2000; Kaufmann, 2006).<sup>1</sup> This is a relatively recent phenomenon. From the dawn of modern polling in the 1930s and on into the 1950s, women identified as Republicans and supported Republican presidential candidates a bit more than men did (Campbell et al., 1980, 493; Ladd, 1997, 120-121). Following a period which gender differences in partisanship were largely absent, the modern partisan gender gap emerged during the late 1970s and early 1980s (Norrande, 1999). While the same pattern wasn't initially present in other advanced industrialized countries (Inglehart, 1977, 228; Norris, 1988, 224), by the 1990s men identified with conservative parties and supported conservative candidates more than women in many of these countries as well (Inglehart and Norris, 2000).

Many different theories of the modern partisan gender gap attribute its emergence to an alleged growing divergence between men's and women's policy preferences. Some scholars highlight social changes, such as greater female labor force participation, liberalized divorce laws, or feminist political socialization among younger generations, that they posit caused growing differences in preferences between the sexes. However, there is little evidence for preference divergence. Roughly since the start of modern polling in the United States, men have consistently expressed more conservative opinions than women on a series of issues, including criminal justice, social welfare, and foreign policy (e.g., Shapiro and Mahajan, 1986; Kaufmann and Petrocik, 1999). Shapiro and Mahajan's meta-analysis finds that gender differences in opinions in the United States were roughly stable from 1964 to 1983, the entire period that they examined and the period when we find that the modern partisan gender gap emerged and showed its largest growth. More recently, Clark (2017) finds that the magnitude of gender preference differences in the General Social Survey

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<sup>1</sup>We follow convention and label these sex differences as a "gender gap," although it is more precisely called a sex gap (Beckwith, 2005).

have been fairly consistent over time.<sup>2</sup>

We argue that ideological party sorting, not any form of preference divergence, was the main mechanism causing the emergence of the partisan gender gap between the 1960s and the 1990s. Since the 1960s, the U.S. public has gradually perceived more and more of the elite-level ideological party polarization occurring during this time (Layman and Carsey, 2006).<sup>3</sup> As this happened, it caused slow, but steady, changes in mass-level party identification as, over many years, people sorted into parties that better matched their policy preferences (e.g., Levendusky, 2009; Abramowitz, 2010; Fiorina, Abrams and Pope, 2011). Because men consistently held more conservative positions than women on several prominent issues, this sorting fueled the modern partisan gender gap's emergence.

We test this explanation using two datasets, each with important, yet different, strengths. The biennial American National Election Studies (ANES) contain detailed questions about respondents' partisan identification, policy preferences, and perceptions of the parties' ideological stances. However, ANES surveys are too small and infrequent to precisely measure variation in the partisan gender gap over time or differences in its size between groups, such as those with more and less awareness of polarization. To remedy this, we assembled the largest dataset ever used to study the partisan gender gap. We pooled individual-level responses to 1,825 Gallup polls that included questions about gender and party identification from 1953 through 2012. While lacking the ANES's variety of political questions, the Gallup dataset gives us much more statistical power to precisely track party identification separately by gender and education (which proxies for awareness of polarization) over time. By leveraging the different strengths of these datasets, we find support for our argument that the modern partisan gender gap emerged largely because party polarization made longstanding gender opinion differences more relevant to partisanship.

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<sup>2</sup>The origins of these consistent differences in the policy preferences of men and women is an important topic, but beyond the scope of this paper. On this, see, for example, Anderson (1997), Sapiro (2003), Huddy, Cassese and Lizotte (2008*a*, 33-35) and Sapiro and Shames (2010).

<sup>3</sup>The pattern of party ideological polarization at the elite and mass level is at least partially driven by liberal and conservative social movements and interest groups increasingly affiliating with the Democrats and Republicans, respectively. This is true on women's rights issues (Mansbridge, 1986; Wolbrecht, 2000) as well as in many other areas (e.g., Karol, 2009).

## 2 Existing Literature

Males' disproportionate support of Ronald Reagan in 1980 sparked scholarly and journalistic interest in the emergence of America's modern partisan gender gap. Since then, scholars have put forth several explanations which rest on the claim that it emerged because of increasing differences between men's and women's policy preferences. One group of explanations claims that women began to prefer a larger welfare state than men because women grew more economically vulnerable. Declining marriage rates and increasing divorce rates are often cited as sources of this increased relative economic vulnerability. In support of this idea, Edlund and Pande (2002) find that the partisan gender gap is larger in states where divorce is more prevalent and that, in panel data, marriage and divorce make men and women more Republican and Democratic, respectively. Similarly, Iversen and Rosenbluth (2006) find that the partisan gender gap is larger in the unmarried population in European countries. Also consistent with this, Box-Steffensmeier, De Boef and Lin (2004) find, with aggregated time-series data, that the U.S. partisan gender gap increases when economic performance wanes and the number of economically vulnerable single women increases.<sup>4</sup>

A second group of explanations contend that growing female labor force participation caused gender policy preferences diverge on some issues. For example, Inglehart and Norris' (2000) Developmental Theory of Gender Realignment claims that women's attitudes on cultural issues, such as freedom of self-expression and gender equality, moved substantially to the left in affluent countries because of their increased labor-market experience and economic independence from men. Along similar lines, Iversen and Rosenbluth (2006, 2011) contend that growing female labor force participation led women to prefer a larger welfare state than men because they needed it to sustain their new economic independence. They cite public sector employment and subsidized daycare as examples of policies that make it easier for females to maintain their careers while raising children. Consistent with this hypothesis, they show that, in European countries, the gender gap in support for public employment and left-leaning political parties is larger among labor force partic-

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<sup>4</sup>Also consistent with an economic vulnerability explanation, several studies show that women tend to perceive the economy more negatively than men do (Miller, 1988; Ladd, 1997; Chaney, Alvarez and Nagler, 1998).

ipants.<sup>5</sup> Carroll (1988, 2006) argues that, by making women less economically and psychologically dependent on men, increased labor market participation allows women to make more independent assessments of their political interests. Huddy, Cassese and Lizotte (2008*a*) call this the “economic autonomy hypothesis.” Consistent with this, Carroll and Manza and Brooks (1998) show, in the United States, that women who work outside the home and in more economically independent professions vote more Democratically.

A third body of work argues that the increasing influence of feminism caused women to become relatively more liberal than men, particularly on social issues. Consistent with this, Inglehart and Norris (2003) find that a substantial portion of the partisan gender gap in wealthy countries is explained by differences in cultural attitudes toward postmaterialism, support for gender equality and the role of government.<sup>6</sup> Also consistent with this idea, Kaufmann (2002) finds that, over time, American women’s attitudes on social issues increasingly correlate with their party identification. Also consistent with this, Conover (1988) and Manza and Brooks (1998) find that women with a feminist consciousness have more liberal policy attitudes and are more likely to identify as Democrats (but see Cook and Wilcox (1991) for a contrary view).

Few of the studies discussed above actually test whether their chosen explanation actually produced any growth in policy preference differences.<sup>7</sup> This is largely because of the sparse amount of data available on issue preferences in the time period that the modern partisan gender gap developed. Instead, these studies generally relate their causal variables to gender differences in partisan identification or vote choice, while assuming that any observed associations occur because the effects flow through divergent policy preferences.<sup>8</sup> A problem here is that the two most comprehensive studies looking directly at American gender differences in policy preferences find little evidence of preference divergence when the partisan gender gap emerged. Shapiro and Mahajan’s (1986) meta-analysis of gender differences on 962 issue questions asked from 1964 through

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<sup>5</sup>However, Deitch (1988) finds little relationship between women’s labor force status and preferences towards government spending over the time period that the modern partisan gender gap emerged in the United States.

<sup>6</sup>However, Inglehart and Norris (2000) find these factors have little explanatory power over the partisan gender gap in the United States.

<sup>7</sup>An exception is Huddy, Cassese and Lizotte (2008*b*, 155), which tests how the gender gap varies across a variety of demographic categories, but is held back by the relatively small sample size of the ANES.

<sup>8</sup>A few of these studies also look at ideological self-placement on a liberal-conservative scale.

1983 shows that gender differences were of a similar magnitude in 1964-1971, 1972-1976, and 1977-1983. Similarly, DiMaggio, Evans and Bryson (1996, 723) find “slim evidence of a growing gender gap” in issues preferences on the ANES and General Social Survey (GSS) between 1972 and 1994. While these studies can be critiqued for aggregating over too many questions, and thus potentially obscuring growth in gender differences on selected items, they conclusively show that men had more conservative policy preferences than women even before the partisan gender gap was present and that overall preference differences did not grow as it emerged.

In what areas do men and women’s policy preferences differ?<sup>9</sup> Shapiro and Mahajan (1986) show that men were significantly more conservative than women on the size of the welfare state and issues related to the use of force, such as national defense and criminal justice policy (see also Anderson, 1997; Sapiro, 2003; Sapiro and Shames, 2010; Barnes and Cassese, 2017; Clark, 2017). Kaufmann and Petrocik (1999) show that men held more conservative social welfare views than women since at least the 1950s.<sup>10</sup> Other studies confirm gender differences on issues related to the use of force (e.g., Smith, 1984; Kaufmann, 2006; Eichenberg and Stoll, 2012), which appear to date back at least to the 1940s (Ladd, 1997, 116).<sup>11</sup>

If men’s and women’s issue preferences have differed in several areas since prior to the 1960s, why didn’t the partisan gender gap form earlier? In the next section, we argue that this is because people did not perceive that the parties were sufficiently differentiated on the issues where men and women disagreed. We are not the first to relate elite party polarization and mass-level sorting to the development of the partisan gender gap. Activists in the 1980s publicized the growing differences between the Democratic and Republican platforms on women’s rights issues, like abortion and the Equal Rights Amendment (ERA), as an explanation for the partisan gender gap (Bonk, 1988;

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<sup>9</sup>Our next two paragraphs review evidence on gender policy preference differences in the overall population. See Barnes and Cassese (2017) on gender policy differences within each party.

<sup>10</sup>While in some political commentary, only opinions explicitly about reproductive freedom and other women’s legal rights are labeled as gender-related issues, many women’s rights organizations plausibly argue that, because of disproportionate poverty among women, poverty and social welfare policy should be considered as another type of gender-rights issue (Sapiro and Canon, 2000, 172).

<sup>11</sup>Findings that other types of policy attitudes have become more related to attitudes about gender equality (e.g., Winter, 2008) are consistent with our argument that people’s attitudes on all major issues became more correlated with partisanship and with each other as people ideologically sorted. As a consequence, some issues where there was a gender preference gap, while not traditionally associated with women’s equality, became thought of as gendered to many people.

Costain, 1988; Wolbrecht, 2000). Yet a problem with that explanation is that men and women usually report similar preferences on these types of issues (Mansbridge, 1985, 1986; Cook, Jelen and Wilcox, 1992; Seltzer, Newman and Leighton, 1997; Anderson, 1997; Sapiro and Conover, 1997; Sapiro, 2003; Sapiro and Shames, 2010; Fiorina, Abrams and Pope, 2011). In addition, the partisan gender gap predates polarization in the national party's positions on these types of issues, which largely occurred after 1976 (Wolbrecht, 2000; Norris, 2003; Kaufmann, 2006), and campaign appeals about traditional women's issues do not seem to increase the partisan gender gap (Hutchings et al., 2004).

Our argument is most similar to, and builds on, Kaufmann and Petrocik's (1999) claim that the partisan gender gap was caused by an increase over time in the influence of social welfare preferences on partisan identification.<sup>12</sup> While we agree with the main thrust of Kaufmann and Petrocik's thesis, we advance their line of argument both theoretically and empirically. We draw on the ideological sorting literature and argue that people increasingly sorted into parties that matched their preferences in all prominent political domains, not just social welfare. We also explain and show that awareness of party polarization is a necessary step in the causal process.<sup>13</sup> Empirically, while Kaufmann and Petrocik (1999) examine only two years of ANES data (1992 and 1996), both after the gender gap had emerged, we employ six decades of ANES and Gallup data.

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<sup>12</sup>Relatedly, Fiorina, Abrams and Pope (2011, 103-104) argue that the Democrats being seen as more dovish in the wake of the Vietnam War made it easier for gender differences in social welfare and use of force preferences to cumulate.

<sup>13</sup>Another precursor to our work is Ladd (1997), who observes that the partisan gender gap is larger among the more educated in the 1980s and 1990s, but does not look at earlier time periods or changes over time, nor does he connect this phenomenon to partisan sorting and perceptions of polarization.

Our argument is consistent with Abramowitz (2010, 81-82). Using ANES data from 1992-2004, he shows demographics more strongly associate with partisanship among the more politically engaged. We build on Abramowitz by showing that the relationship exists in more datasets, showing how it developed over time, and providing more evidence that it is driven by preference-based sorting.

Burden (2008) finds that the partisan gender gap is substantially attenuated, particularly among the most politically aware, when people are asked about which party they feel a part of, instead of which party they think of themselves as being. To the extent that priming thinking instead of feeling increases the weight that people put on issue preferences when formulating their partisan identification, Burden's finding is consistent with our theory.

### 3 Theory and Hypotheses

Since the 1960s, Democratic and Republican politicians more consistently took liberal and conservative positions, respectively, on a wide range of prominent national political issues, including the size of the welfare state, crime, civil rights, and military aggressiveness (Carmines and Stimson, 1989; Noel, 2013; McCarty, Poole and Rosenthal, 2016). Over time, some members of the public noticed that their issue preferences were increasingly inconsistent with their party loyalties (Layman and Carsey, 2006). Slowly, some of these people adjusted their preferences to match their partisanship, while others did the opposite: sorting into a party that better matched their policy preferences (Layman and Carsey, 2006; Levendusky, 2009; Abramowitz, 2010; Fiorina, Abrams and Pope, 2011).

We theorize that this ideological sorting led to the emergence of the modern partisan gender gap. As we discussed in the previous section, men consistently held more conservative policy views than women in major issue areas, even prior to the emergence of the partisan gender gap. This meant that there were more Democratic men than women with conservative issue preferences and more Republican women than men with liberal issue preferences. Thus, ideological sorting led relatively more men than women to move from the Democrats to the Republicans and relatively more women than men to move from the Republicans to the Democrats. Because the partisan gender gap was emerging when Republican identification was near an all-time low following Watergate, there were more out-of-sync conservative Democrats than out-of-sync liberal Republicans. Thus, an extra implication of our theory is that the partisan gender gap developed more because of change in men's than women's partisanship, which is consistent with the findings of Wirls (1986), Kaufmann and Petrocik (1999), and Norrander (1999), among others.

We expect that the partisan gender gap gradually emerged in response to awareness of increasing party polarization. It is well-established that party identification is a sturdy attitude, that only responds to major political change, and even then does so slowly (Campbell et al., 1980; Jennings and Niemi, 1981; Green, Palmquist and Schickler, 2002). In the short term, people rarely leave a party that is out of sync with their issue preferences. Instead, some adjust their issue preferences to

match their partisanship (Lenz, 2012), particularly when the perceived differences between ideology and partisanship are not large (Levendusky, 2009). Moreover, McCarty, Poole and Rosenthal (2016) show that elite polarization occurred slowly at first, before becoming more rapid in the late 1970s. And even once elite polarization occurred, it may have taken some time for people to be aware of it. All of this indicates that the type of sorting we are concerned with—changing one’s partisanship to match issue preferences—does happen, but only gradually over many years (Levendusky, 2009; Achen and Bartels, 2016).

Our theory predicts that we should observe a partisan gender gap first among those who perceive the most polarization between the two parties. Previous work shows that education and other measures of political engagement are positively associated with knowledge of elite political positions, especially when those positions change over time (e.g., Converse, 1964; Zaller, 1992, 1996). Thus, if the partisan gender gap is ultimately driven by people noticing the increasing ideological polarization of the two parties, it should emerge earlier and be larger among those with more education.

We can summarize our main argument with three main hypotheses, the second of which has two parts:

(H1) The partisan gender gap should emerge slowly as the parties become more polarized.

(H2) The partisan gender gap is larger among those that perceive more political polarization.

(H2a) The partisan gender gap should be larger among the more educated.

(H2b) Because the relationship between education and the partisan gender gap flows through perceptions of polarization, that relationship should shrink when one controls for those perceptions.

(H3) At the same time that the partisan gender gap emerged, the association between individuals’ issue preferences and partisanship should have increased, especially among college educated men and women, who have (according to H2a) the largest partisan gender gap.

## 4 Data

We test these hypotheses using two datasets, each with their own different strengths: the ANES and a new individual-level dataset of pooled responses to Gallup polls from 1953 through 2012. As the ANES is used widely in political science, we will not describe it in detail here, but simply highlight its advantages for our project. Since 1970, the ANES asked respondents about their perceptions of the positions of the Democratic and Republican parties in a variety of issue domains. This allows us to directly relate perceptions of polarization to the partisan gender gap. The ANES also contains detailed policy preference questions on a variety of areas during the decades when this gender gap arose. This allows us to study how the gender gap in issue preferences changed before and after the emergence of the modern partisan gender gap.

Much of the existing literature on the partisan gender gap examines only the ANES (e.g., Wirls, 1986; Chaney, Alvarez and Nagler, 1998; Kaufmann and Petrocik, 1999; Norrander, 1999; Edlund and Pande, 2002; Kaufmann, 2002; Norris, 2003; Huddy, Cassese and Lizotte, 2008*a*; Kaufmann, 2006; Sapiro and Shames, 2010), even though it only contains somewhere between 1,000 and 3,000 respondents, sampled usually every two years since 1952. Several studies use the GSS instead (e.g., Wirls, 1986; Deitch, 1988), which has a similar sample size and has been conducted at similar intervals since 1972. To measure exactly when the partisan gender gap emerged, and among whom, we need bigger samples. This is particularly true when examining whether the gap's size depends on attributes like education, because comparing the gap between subpopulations requires a sufficient sample in each of them.

To improve on the ANES and GSS's statistical power, we pooled frequent Gallup surveys to assembled the largest dataset ever used to study the partisan gender gap. While some have used aggregate commercial survey data previously to model the dynamics of men's and women's partisanship (Box-Steffensmeier, De Boef and Lin, 2004) and presidential approval (Clarke et al., 2005), we are aware of no previous studies that has used individual-level data to do so. In Section 7.1 in the Appendix, we describe our collection of individual-level responses from every poll conducted by the Gallup Organization from 1953 through 2012 that (1) asked about presidential approval, party

identification, and/or ideology and (2) is contained in the Roper Center iPOLL database.<sup>14</sup> In the dataset, there are 2,250,703 observations from 1,825 surveys that ask a nationally representative sample about their gender and partisan identification. Because at least 13, and often substantially more, Gallup polls are available in every year our dataset covers, the pooled dataset contains tens of thousands of respondents per year. This gives us greater statistical power to precisely measure how the partisan gender gap changes over time and to examine subgroups (e.g. college graduates) separately, while maintaining adequate sample sizes.<sup>15</sup>

There are also drawbacks to our Gallup dataset. Many demographic characteristics and political attitudes that we would like to observe are asked sporadically, if at all. Given the substantial cost of merging each additional variable, and the limited usefulness of merging a variable only asked in, for instance, 2 or 3 surveys, we only merged variables into the pooled dataset that were asked relatively consistently over time.<sup>16</sup> The only attitudes probed consistently enough to make them useful to merge together were presidential approval, party identification, and, since 1992, ideological self-placement. Thus, we are unable to examine specific issue preferences over time using Gallup data.

Gallup data also do not contain direct questions about perceptions of the ideological positions of the parties or several other types of questions used to measure political awareness in the ANES.<sup>17</sup> We use education as a proxy for respondents' awareness of changing elite party positions and ability to understand the consequences of those positions for their own party affiliation. Because education is strongly correlated with media exposure and overall political sophistication (Fiske, Lau and Smith, 1990; Price and Zaller, 1993; Delli Carpini and Keeter, 1996), it is often used as a proxy for these types of attributes in public opinion scholarship (e.g., Brady and Sniderman,

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<sup>14</sup>We start in 1953 because Dwight Eisenhower's was the first presidency during which Gallup used probability sampling exclusively (Berinsky, 2006).

<sup>15</sup>Our Gallup series begins well before the emergence of the modern gender gap, unlike the aggregate time series of CBS/*New York Times* polls used by Box-Steffensmeier, De Boef and Lin (2004), which starts in 1979.

<sup>16</sup>See Section 7.1 in the Appendix for a description of the variables we collected. Of course, a large number of questions asked in Gallup surveys were only asked with the same question wording once. In these cases, merging the question into our combined dataset provides no advantage over just analyzing that poll separately.

<sup>17</sup>Other measures of overall political awareness often asked in the ANES, but not Gallup, are responses to questions about basic political facts (Price and Zaller, 1993) and the interviewer's assessment of the respondent's level of political awareness (Bartels, 1996).

1985; Carmines and Stimson, 1989; Zaller, 1994; Berinsky, 2009; Hayes and Guardino, 2013). For example, Carmines and Stimson (1989, ch. 5) argue that the politically “sophisticated” are more likely to notice long term changes in party positions, and use education, in addition to political information questions, to measure political sophistication.

Gallup asks about party identification in a slightly different manner than the ANES, GSS, and CBS/*New York Times* polls. Gallup asks “In politics, as of today, do you consider yourself a Republican, Democrat or Independent?”<sup>18</sup> Abramson and Ostrom (1991) argue that the Gallup question, which is also used in Erikson, MacKuen and Stimson’s (2002) analysis of “macropartisanship,” introduces more short-term political and economic considerations than the standard party identification question. Yet while important, these differences should not be overstated. ANES party identification and Gallup party identification are substantially correlated both within individuals and in their aggregate movements over time (MacKuen, Erikson and Stimson, 1992; Kaufmann and Petrocik, 1999, 868-870).

A related complication is that Gallup did not always follow-up with Independents and ask whether they lean towards a party. They did so occasionally in the 1950s, almost never in the 1960s and 1970s, sometimes in the 1980s, and then regularly from the 1990s on. As we cannot consistently observe Independent leaners in the Gallup data, we code them as Independents in our main analysis. Because men are more likely than women to label themselves as Independent leaners rather than partisans (Norrander, 1999), we check the robustness of results to including leaners as partisans on the subset of surveys that ask about leanings (see Section 7.5 in the Appendix).

A final issue is that Gallup changed from using in-person surveys to phone surveys over time. When both modes were used during the late 1980s and early 1990s, it showed that phone respondents tend to be more Republican than in-person respondents (Green, Palmquist and Schickler, 2002). Thus, we need to take care when looking at long run changes in the partisan gender gap to account for any differences caused by this change in survey mode.

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<sup>18</sup>Gallup also occasionally omits “as of” from the question.

## 5 Results

### 5.1 The Gradual Increase in the Partisan Gender Gap

Hypothesis #1 (H1) predicts that the partisan gender gap emerged gradually as people became aware of the increasingly large differences between Democratic and Republican elites' policy positions. What can be observed in surveys like ANES and GSS is that the genders start moving toward the modern pattern sometime in the 1960s, with the modern party identification gap first reaching statistical significance in the 1970s, and becoming consistently significant in every ANES survey by the late 1980s (e.g., Miller, 1988; Kaufmann and Petrocik, 1999; Norrander, 1997, 1999; Sapiro and Shames, 2010). Given the sample size of surveys like the ANES and GSS, it is neither possible to tell whether the small partisan gender gaps in some years in the 1960s, 1970s and 1980s is evidence that a partisan gender gap is really present in the population, nor whether the often large year-to-year fluctuations reflect real changes in the partisan gender gap or random sampling variation. Consequentially, existing studies that use ANES or GSS data give varying answers as to when, and how suddenly or gradually, the partisan gender gap first emerged. Kaufmann and Petrocik (1999) identify 1964 as the origin of the partisan gender gap because a higher percentage of women have identified as Democrats than men in presidential ANES surveys since then. However, Norrander (1999) notes that this is partially an artifact of men being more likely to initially identify as Independents; she shows that throughout much of the 1970s women were more likely to identify both as Democrats and Republicans than men. When leaners are classified as partisans, there is a statistically significant partisan gender gap in the ANES of about four percentage points that first appears in 1972 and 1974, largely disappears in 1976 and 1978, only to reemerge again in 1980 (Norrander, 1997).

Another consequence of small samples sizes is that the literature often faces a tension between defining the partisan gender gap in terms of substantive or statistical significance. A focus on statistical significance risks missing the partisan gender gap's emergence, as a modest gender gap will not be statistically distinguishable from zero in a survey with the ANES or GSS's sample size. But focusing instead on point estimates risks over-fitting to explain random variation. Our

pooled Gallup dataset alleviates this problem because it combines surveys that are frequent and numerous enough, and thus have a sufficiently large combined sample size, that we can more precisely identify when the partisan gender gap first emerged and its size over time. The smaller standard errors on our estimates of the gap allow us to focus more on its substantive magnitude, while still remaining cognizant of the potential for sampling error.

To analyze the partisan gender gap over time with the Gallup dataset, we construct the sample-weighted average partisanship of men and women, respectively, within each survey  $s$  in our sample.<sup>19</sup> We take a non-parametric approach in which each gender's partisanship at time  $t$  is estimated by taking a weighted average of surveys that occur in close proximity to time  $t$ . A key consideration when constructing these weighted averages is determining how much weight is given to a specific survey  $s$  based on the proximity of the date of the survey, labeled  $t_s$ , to time  $t$ . We define  $t_s$  as the midpoint of when survey  $s$  was in the field. Based on the results of a leave-one-out cross-validation procedure, we use a Epanechnikov kernel function with a bandwidth of 100 days to construct these weighted averages throughout our analysis.<sup>20</sup>

The top panel of Figure 1 presents the evolution of partisanship by gender from 1953 through 2012. Several trends stand out from this broad overview. In the 1950s, American women were slightly more likely than men to identify with the Republicans (Anderson, 1997; Ladd, 1997; Sapiro and Canon, 2000, 171). However, the overall partisanship of men and women was quite similar from the early 1960s through the mid 1970s. While both men and women became more Republican from the late 1970s to the early 1990s, men did so at a faster rate. As a result, a substantial partisan gender gap emerged over this time period that remains in place through 2012.

One can see these patterns most clearly in the bottom panel of Figure 1, which graphs the partisan gender gap over time. There was not much of a gap before the late 1970s. The most rapid change in the gap occurred between 1977 and 1980, when it went from roughly zero to about four percentage points.<sup>21</sup> From 1980 until 1997, the gap oscillated between about three and

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<sup>19</sup>Figure A.7 in Appendix Section 7.8 plots both of these quantities for each  $s$  in our sample.

<sup>20</sup>See Section 7.2 in the Appendix for more details.

<sup>21</sup>This contrasts with Box-Steffensmeier, De Boef and Lin (2004), who find a reverse partisan gender gap in CBS/*New York Times* polls in early 1979, followed by a rapid increase in 1979 and 1980 in the percentage of women versus men who identify as Democrats.

six percentage points on in-person Gallup surveys, staying consistently significantly different than zero, with the exception of a couple of instances where relatively less data caused the standard errors to spike.

The partisan gender gap remained fairly stable in the 1990s and 2000s. In the 1990s, Gallup phased out in-person political polling in favor of phone polling. To ensure that mode effects are not confused with real opinion change, Figure 1 graphs the in-person and phone poll trends separately. There are two mode differences. Both sexes are more Republican on phone surveys than on in-person surveys, producing a mode-driven overall shift toward the Republicans in the 1990s. But our real concern is the gender gap, not the overall partisanship level. The second mode difference is that the partisan gender gap was about two points higher in phone surveys than in-person surveys conducted in the same time period. Thus, the apparent growth in the partisan gender gap in the 1990s appears to come largely from the mode switch. In the 2000s, when Gallup's transition to phone was complete, the gap remained steady at about seven percentage points. Because in-person polls were used during the bulk of the gender gap's growth, followed by stability in the late 1990s and 2000s, we simplify some of our subsequent analyses of the Gallup data by using only the in-person polls at little explanatory cost.<sup>22</sup>

Figure 2 illustrates the advantage of the pooled Gallup dataset's large sample size for testing H1 by comparing the gender gap estimates in the Gallup data with every ANES and GSS survey during this period. When viewed together, all three series appear to follow the same trend. The partisan gender gap from in-person Gallup surveys is within the ANES's 95% confidence interval 20 out of 22 times and the GSS's 95% confidence interval 19 out of 21 times. Likewise, the partisan gender gap estimate from phone Gallup surveys is within the ANES's 95% confidence interval 10 out of 10 times and the GSS's 95% confidence interval 13 of 14 times. This is roughly the proportion of the Gallup estimates that one would expect to fall outside the 95% confidence intervals due to sampling error. This reassuringly suggests that, despite question wording differences, all three

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<sup>22</sup>As we noted earlier, leaners are treated as Independents because the question about which party Independents lean toward was not asked in all surveys. Section 7.5 in the Appendix examines how results change if we classify Independent leaners as partisans in the subset of Gallup surveys that followed up by asking Independents about their leanings. Results show that our key findings on when the partisan gender gap emerged and the presence of educational heterogeneity in the partisan gender gap hold when we classify Independent leaners as partisans.

Figure 1: Smoothed Partisanship Level by Gender in Gallup Surveys

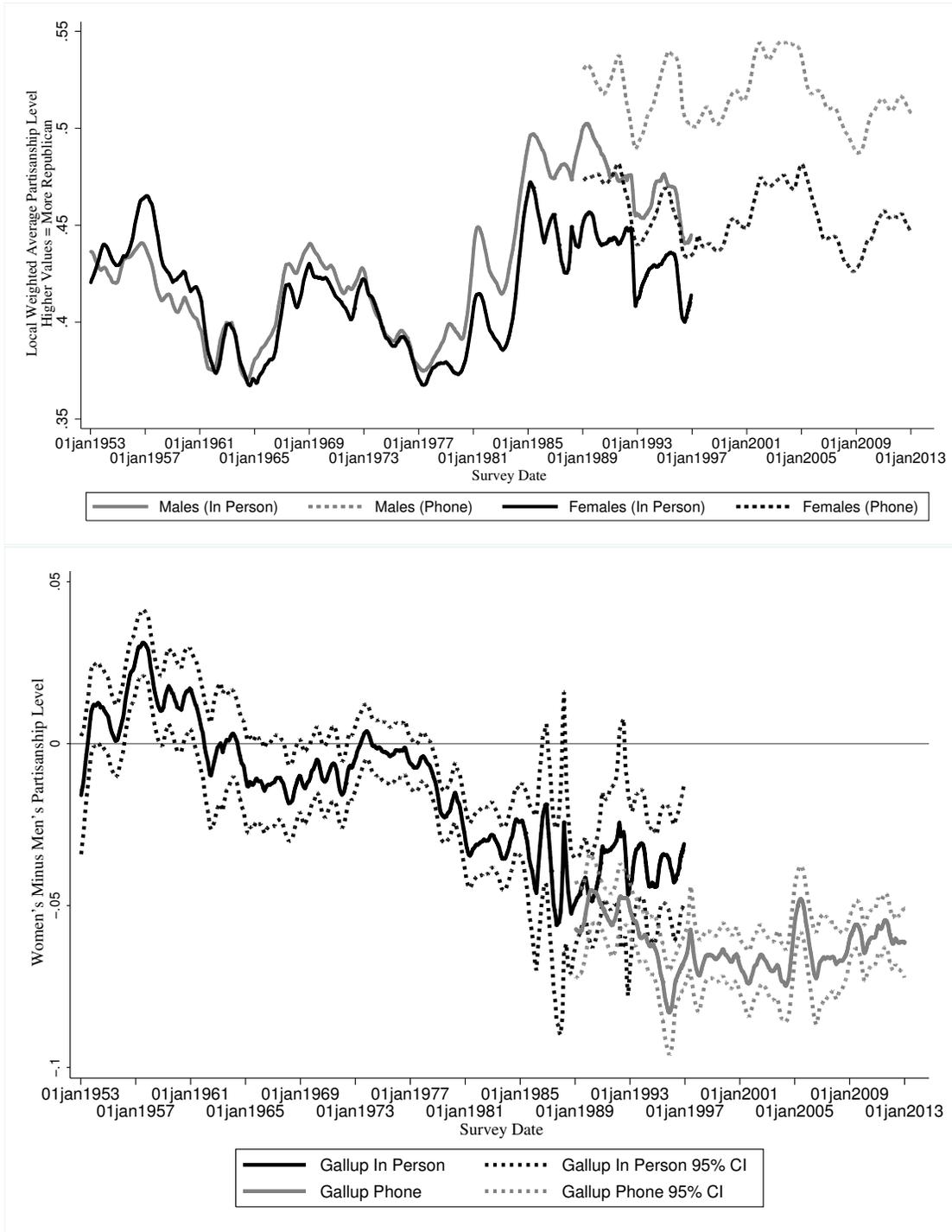
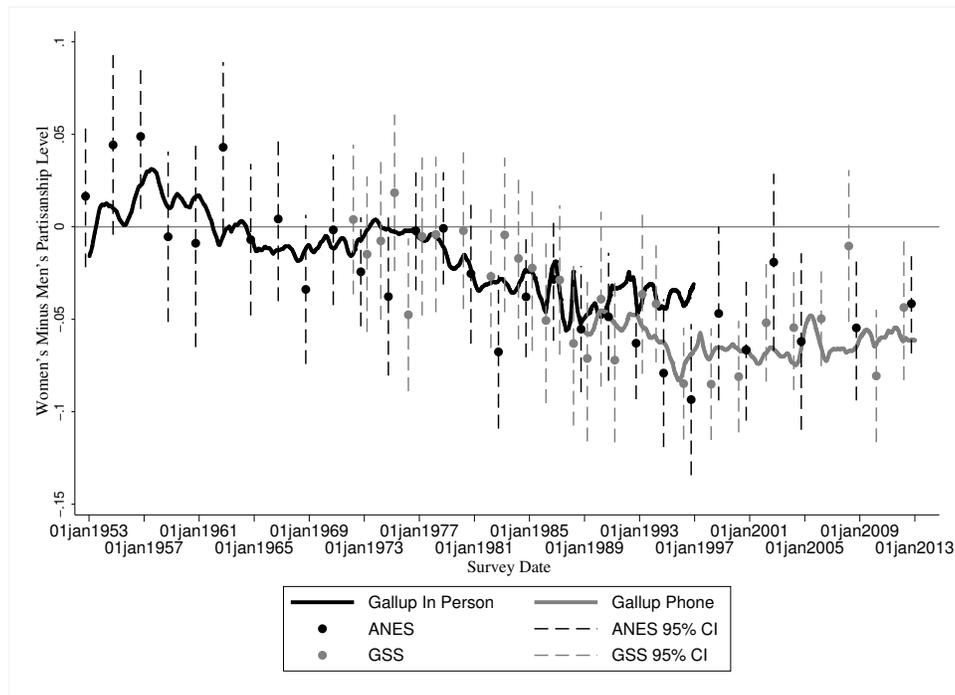


Figure 2: Comparing Partisan Gender Gap in Gallup, ANES, and GSS Surveys



surveys capture a similar construct. The main difference is that the smaller sample sizes in the ANES and GSS produce much larger confidence intervals, which make them unable to detect a significant partisan gender gap in some years where it is present. For example, ANES and Gallup generate nearly identical point estimates of the partisan gender gap in the late 1980s, but we can only reject the null of no partisan gender gap using Gallup data. Figure 2 shows that, consistent with H1, the development of the partisan gender gap is a smoother process than one might conclude from the ANES or GSS. Several surges and swoons in the gap's size in the ANES or GSS, which one might be tempted to imbue with political importance if one considered these surveys alone, now appear to be mere sampling variation around the gradual trend.<sup>23</sup>

Whether ultimately caused by partisan sorting or some other mechanism, it is possible that the

<sup>23</sup>For instance, it appears in the ANES that, after disappearing in 1958 and 1960, the old reverse gender gap re-emerged for the last time in 1962. Years such as 1968 and 1974 stand out as points when the modern gap first emerged at notable sizes (and marginal statistical significance). More recently, 1982, 1994, and 1996 stand out for particularly large gender gaps. It appears in the GSS that the modern gender gap first emerged in 1976 (in contrast to the ANES), and then dissipated until emerging again in the mid-1980s, then temporarily shrank in 1993-1994. But all this appears to be largely sampling variation.

partisan gender gap’s growth was driven by specific political events. Using ANES data, Carmines and Stimson (1989, ch. 6) argue that 1964 was a “critical moment” that triggered a lot of immediate mass-level party sorting based on racial policy preferences, followed by slower party sorting in the years after. Kaufmann and Petrocik (1999) identify the 1964 and 1980 presidential campaigns as instances where the Republican Party’s positions on social welfare policy moved substantially to the right, raising the salience of those positions and thus causing the gap’s growth. If specific events or years led to sudden growth in the gender gap, it would contradict H1.

However, Figure A.2 in Appendix Section 7.3 illustrates that both the magnitude of, and trend in, the partisan gender gap appear to be very similar before and after these presidential campaigns. In that section, we also apply a formal statistical test and find no evidence that either the level or the slope of the partisan gender gap changed before or after the 1964 and 1980 presidential campaigns. The evolution of the partisan gender gap appears to be similar before and after 11 of 13 presidential campaigns between 1960 and 2008, with the exceptions being 1976 and 1996. Thus, consistent with H1, our large Gallup dataset reveals that the sorting leading to the partisan gender gap was a slow process. In Carmines and Stimson’s (1989, 139) typology, we find that the sorting process follows more of a “secular realignment” than a “dynamic growth” pattern.<sup>24</sup>

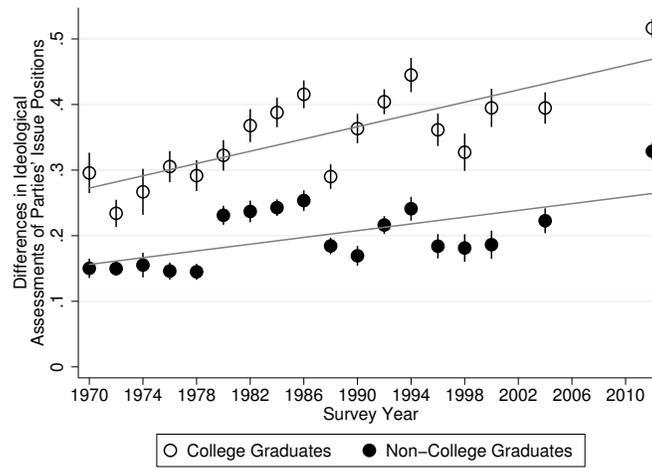
## **5.2 The Gender Gap is Associated with Knowledge of Elite Party Polarization**

Our theory predicts that the partisan gender gap should be larger among those who are more educated, because such individuals are more likely to be aware of increased polarization. This prediction is supported by the data presented in Figure 3, which compares assessments of polarization in the ANES over time among college graduates and non-college graduates. The dependent variable in this graph is constructed using assessments of the conservatism of the Democratic and Republican party’s issue positions on all available issues, including domestic welfare spending, law and order, racial policy, and gender-related issues from 1970, when these questions were first asked

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<sup>24</sup>Of course, in addition to using a larger dataset, we are examining a different question than Carmines and Stimson (1989) did. They looked only at sorting on racial issues, while our focus is sorting on a range of issues.

Figure 3: Assessments of Polarization in the Parties' Issue Positions (ANES 1970 - 2000, 2004, 2012)



Notes: The dependent variable is the mean difference in respondents' assessments of the average conservatism of the Republican Party's and Democratic Party's positions on all available issues. A value of zero corresponds to an assessment that the issue positions of both parties were equally conservative, while a value of one corresponds to an assessment that the Republican Party's positions were maximally conservative and the Democratic Party's positions were maximally liberal. Black vertical lines indicate the 95% confidence interval on the point estimate in a given year. Grey lines indicate linear trends on point estimates over time.

in the ANES, through 2012. A respondent's assessment of a party's issue position is recoded so that "1" equals the maximally conservative position and "0" equals the maximally liberal position. We calculate a respondent's assessment of polarization on a given issue by taking the difference between their assessment of the Republican and the Democratic Party's positions on the issue. We aggregate over all of the issue positions asked in a given survey to construct a respondent's overall assessment of polarization.<sup>25</sup> The white and black circles represents the average overall assessment of polarization by college graduates and non-college graduates, respectively, in a given year, and the grey lines show the linear trends over time.

Two patterns emerge in Figure 3. In every year, college graduates assess the parties to be further apart ideologically than those with less education do. Moreover, the gap between the college and non-college educated's assessments of polarization increases over time.

This leads to Hypothesis #2, which has two parts. The first, H2a, predicts that these differences

<sup>25</sup>Section 7.4 in the Appendix shows similar patterns when we account for changes in the issues that are surveyed in different years.

in perceptions of polarization will lead to differential sorting, which will cause the more educated to exhibit a larger partisan gender gap. The pooled Gallup dataset is useful for testing H2a because its large size allows us to more precisely examine differences in the gender gap between those with more and less education.

The top row of Figure 4 compares gender differences in the partisanship of college graduates (left) to non-college graduates (right). We restrict our analysis to in-person surveys in this subsection to ensure the results aren't driven by changes in survey mode. The slow and steady growth in the aggregate partisan gender gap displayed in Figure 1 masks large differences across education levels. Among college graduates, in the 1950s, male and female partisanship was similar. Yet men were more resistant than women to the pro-Democratic macropartisanship trends of the 1960s and 1970s, and women were almost entirely unmoved by the pro-Republican macropartisanship trend in the 1980s. On the other hand, for those without a college degree, there is no sign of a partisan gender gap before 1980. During the 1980s, both sexes were influenced by the overall pro-Republican macropartisanship trend. But men embraced it more, creating a significant gender gap, albeit one that was still smaller than among college graduates.

The graph in the bottom left quadrant of Figure 4 compares the size of the partisan gender gap for college graduates and non-college graduates. The right graph on the bottom row of Figure 4 plots the difference in the size of the partisan gender gap between these two groups. The solid black line shows the difference in the size of the estimated gap, with the dotted lines bounding the 95% confidence interval on that estimate. Consistent with H2a, the partisan gender gap was significantly larger among college graduates than among non-college graduates from the early 1970s onward.

Figure 5 shows why the pooled Gallup dataset was necessary to detect differences in the size of the partisan gender gap across education levels over time. The solid black line shows the difference in the size of this gap between college and non-college educated in the pooled Gallup dataset. The same difference in the ANES and GSS is shown with the dark and light dots, respectively, with the

Figure 4: Partisanship by Gender and Education (Gallup)

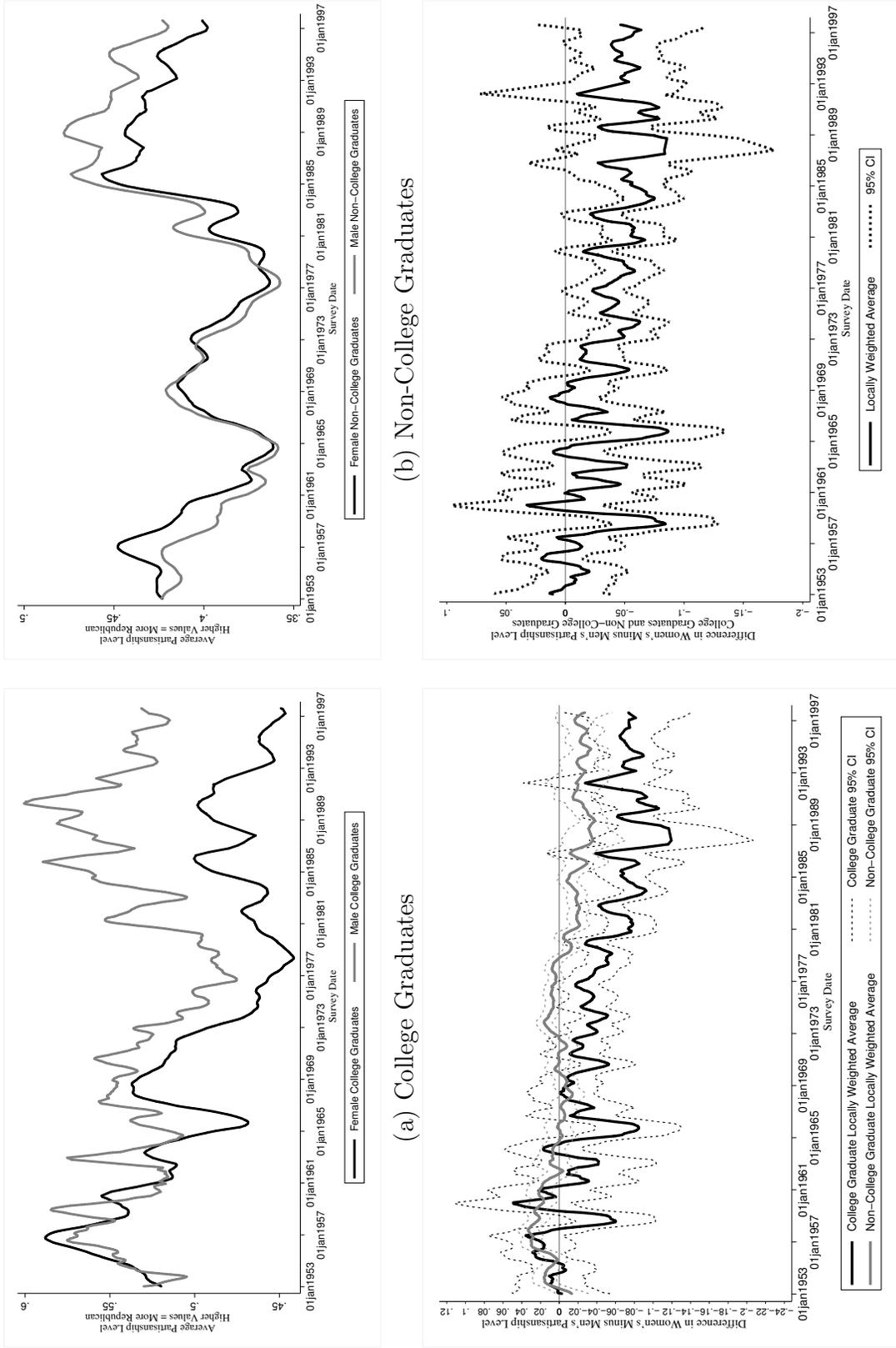
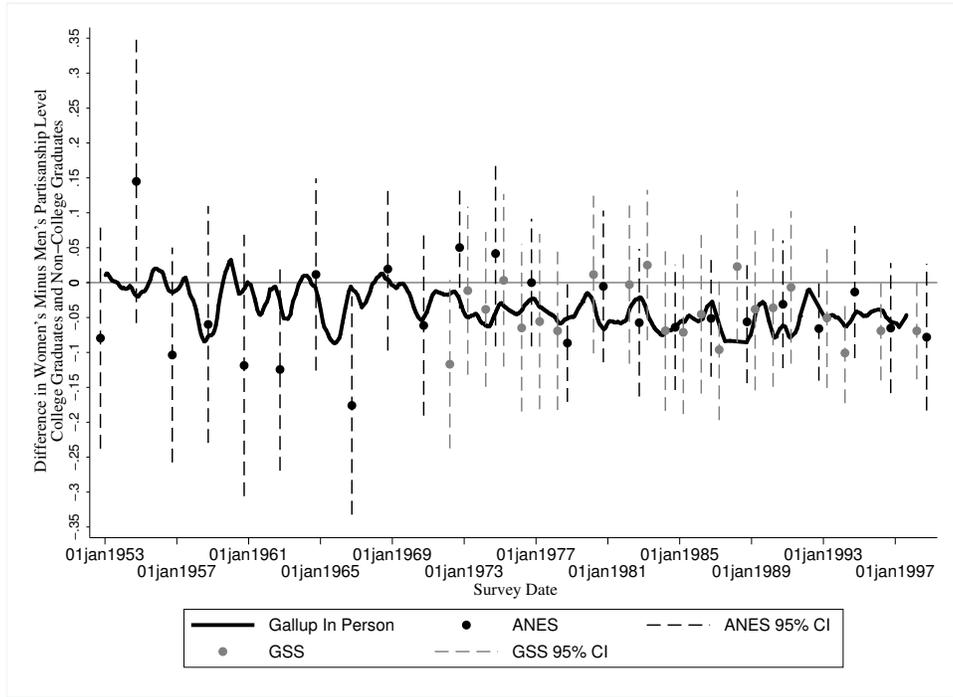


Figure 5: Comparing Difference in Partisan Gender Gap by Educational Attainment in Gallup, ANES, GSS Surveys



dashed lines showing the 95% confidence intervals.<sup>26</sup> The smaller sample sizes in the ANES and GSS produce confidence intervals that are too large to detect this heterogeneity in the partisan gender gap. One can detect an educational difference by pooling many ANES surveys together, as we do later in this section, but there isn't enough sample size to show that these differences have been fairly consistent since the early 1970s.<sup>27</sup>

While we contend that the partisan gender gap is larger among college graduates because of their greater awareness of elite polarization, there are many of variables, such as race, age, region, and economic circumstance, that are associated with being a college graduate. One of these other variables could be driving the relationship between education and the size of the partisan gender

<sup>26</sup>As in Figure 4, the Gallup data includes only in-person interviews and uses the same Epanechnikov kernel smoothing function.

<sup>27</sup>Also, similar to Figure 2, we are reassured to see that the Gallup point estimate is within the 95% confidence interval of every ANES and GSS survey except the 1966 ANES. This pattern suggests again that all three surveys are measuring the same underlying pattern. The difference is simply the smaller sample sizes, which produce greater confidence intervals and year-to-year variation in the ANES and GSS.

gap.<sup>28</sup> The sample size in the ANES and GSS is a problem here is well. Those surveys do not have enough statistical power to simultaneously control for other interactions with gender to see whether it is really education that produces a difference in the size of the partisan gender gap instead of another correlated attribute. Fortunately, our pooled Gallup dataset does have enough sample size to include those controls.

To control for other variables associated with education, we estimate Equation 1. Our key parameter of interest is  $\theta$ , which measures the difference in size of the partisan gender gap among those who have and have not graduated from college. We control for a vector of other characteristics that may influence the partisanship of respondent  $i$  on survey  $s$ , labeled  $X_{s,i}$ , as well as the interaction between these covariates and a female indicator, labeled  $Female_{s,i}$  and a college-graduate indicator,  $College_{s,i}$ . Finally, we include survey fixed effects, labeled  $\alpha_s$ , to account for features of the political environment that affect the overall partisanship of the population at time  $t_s$ . If  $\theta$  is capturing the effect of awareness of polarization, we expect our estimate of  $\theta$  to be relatively unaffected by the inclusion of these controls.

$$Prtnshp_{s,i} = \alpha_s + \beta X_{s,i} + (\gamma + \delta X_{s,i})Female_{s,i} + (\lambda + \kappa X_{s,i})College_{s,i} + \theta Female_{s,i} College_{s,i} + \epsilon_{s,i} \quad (1)$$

Table 1 shows how estimates of educational differences in the partisan gender gap vary over time. Each cell entry presents an estimate of  $\theta$  from a separate regression over the specified time period and including the specified controls. To give each regression more statistical power to accommodate the controls, we group the data into 4-year periods.

We face a trade-off between adding more controls and not discarding observations in the Gallup data. Because Gallup did not always ask all of the demographics that we would like to use, we are forced to drop surveys as we increase the number of controls. The regressions reported in column 1 include all surveys. In columns 2 and 3, we use all surveys that contain a set of “baseline controls”

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<sup>28</sup>Tables A.14-A.17 in Appendix Section 7.8 show the bivariate relationship between the size of the partisan gender gap and all of our demographic variables.

that are available in nearly every Gallup poll: race, decade of birth, and region. Columns 4-6 also control for income, which requires dropping surveys where income was not asked, such as all those prior to the late 1950s. Finally, columns 7-9 also include marital and employment status, which requires also dropping surveys where employment was not asked, which includes all those prior to the late 1970s.

Consistent with H2a, we generally find that educational differences in the size of the partisan gender gap attenuate only slightly when controls are included. Comparing columns 2 and 3 shows that educational differences decline only slightly when we control for the respondent's race, region of residence, and decade of birth. Comparing columns 5 and 6 shows that adding income as an additional control has little effect on the point estimates. Likewise, comparing columns 8 and 9 shows that our estimates of  $\theta$  are roughly the same when we control for employment and marital status. The robustness of our estimates of  $\theta$  to controls is consistent with our claim that college graduation status proxies for awareness of polarization, rather than other attributes like income or labor market experience.<sup>29</sup>

Next, we turn our attention to H2b, which predicts that, because the relationship between education and the gender gap flows through perceptions of polarization, that relationship should shrink when those perceptions are controlled for. To check this, we turn back to the ANES because it contains questions about perceptions of polarization. Table 2 shows a series of regression models using pooled ANES data from all of the years in which it contained questions about the parties' issue positions. The dependent variable in this analysis is partisan identification with leaners coded as Independents, in order to make these results as comparable as possible with the Gallup results.<sup>30</sup> To start, column 1 simply shows, consistent with H2a, that the partisan gender gap is roughly twice as large among college graduates than among non-college graduates in the ANES over these years.

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<sup>29</sup>Table A.12 in the Appendix Section 7.4 shows that including leaners as partisans slightly increases the difference in partisanship between college and non-college graduates in the Gallup data. Another way to check the robustness of our findings is, rather than control for a series of variables, to estimate the gender gap separately in each of these different groups over time in the Gallup data. We do this and find nothing to substantially contradict our argument. The results are in Tables A.14 through A.17 in Section 7.8 in the Appendix.

<sup>30</sup>Table A.18 reports the results of the same regressions with leaners coded as partisans, which show similar findings.

Table 1: Effects of Controls on Educational Heterogeneity in the Partisan Gender Gap (Gallup)

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
<b>Controls:</b>									
Baseline Controls	No	No	Yes	No	Yes	Yes	No	Yes	Yes
+ Income	No	No	No	No	No	Yes	No	No	Yes
+ Employment & Marital Status	No	No	No	No	No	No	No	No	Yes
<b>Year of Survey:</b>									
1953 - 1956	0.000 (0.010) N = 108326	0.001 (0.010) N = 105449	0.007 (0.009)						
1957 - 1960	-0.014 (0.010) N = 107234	-0.015 (0.010) N = 98816	-0.014 (0.010)	0.060 (0.049)	0.017 (0.051)	0.025 (0.053) N = 5045			
1961 - 1964	-0.011 (0.009) N = 147305	-0.014 (0.010) N = 138148	-0.015 (0.009)	-0.015 (0.010)	-0.016 (0.009)	-0.013 (0.010) N = 134181			
1965 - 1968	-0.040 (0.008) N = 141313	-0.037 (0.009) N = 139290	-0.038 (0.008)	-0.039 (0.009)	-0.041 (0.009)	-0.036 (0.009) N = 132281			
1969 - 1972	-0.021 (0.008) N = 113332	-0.019 (0.008) N = 112176	-0.015 (0.008)	-0.018 (0.008)	-0.015 (0.008)	-0.014 (0.009) N = 110432			
1973 - 1976	-0.044 (0.007) N = 135406	-0.043 (0.007) N = 132560	-0.034 (0.007)	-0.042 (0.007)	-0.033 (0.007)	-0.031 (0.007) N = 126698			
1977 - 1980	-0.045 (0.007) N = 150859	-0.045 (0.007) N = 149332	-0.038 (0.007)	-0.045 (0.007)	-0.038 (0.007)	-0.037 (0.007) N = 146539	-0.045 (0.009)	-0.039 (0.009) N = 70881	-0.043 (0.010)
1981 - 1984	-0.052 (0.007) N = 124148	-0.052 (0.007) N = 122030	-0.046 (0.007)	-0.053 (0.007)	-0.047 (0.007)	-0.049 (0.007) N = 119028	-0.056 (0.007)	-0.049 (0.007) N = 107912	-0.050 (0.008)
1985 - 1988	-0.058 (0.013) N = 31354	-0.058 (0.013) N = 31267	-0.045 (0.013)	-0.057 (0.013)	-0.044 (0.013)	-0.045 (0.014) N = 30857	-0.058 (0.013)	-0.044 (0.013) N = 30245	-0.043 (0.015)
1989 - 1992	-0.048 (0.012) N = 36024	-0.048 (0.012) N = 35878	-0.047 (0.012)	-0.049 (0.012)	-0.047 (0.012)	-0.049 (0.013) N = 35204	-0.051 (0.013)	-0.051 (0.013) N = 32165	-0.054 (0.014)
1993 - 1996	-0.048 (0.010) N = 45247	-0.046 (0.010) N = 44967	-0.046 (0.010)	-0.046 (0.010)	-0.046 (0.010)	-0.044 (0.011) N = 44084	-0.044 (0.010)	-0.044 (0.010) N = 43507	-0.039 (0.011)

Notes: Each cell reports the estimate and standard error of  $\theta$  when estimating equation 1 over the specified time period with the specified set of controls. Baseline controls are an African-American indicator, decade of birth indicators, and region of residence indicators. Income controls include indicators for being in the top 20th and 50th percentiles of household income in a survey's sample. Employment and marital status controls include indicators for being married, being employed full time, and being employed part time. Observations are weighted by their sample weight. Robust standard errors are reported in parentheses.

Table 2: Education, Assessments of Party Polarization, and Partisan Identification (ANES, 1970 - 2000, 2004, 2012)

	(1)	(2)	(3)
Female	-0.037 (0.005)	-0.024 (0.006)	-0.020 (0.006)
College Graduate	0.098 (0.008)		0.085 (0.008)
Female X College Graduate	-0.036 (0.012)		-0.022 (0.012)
Polarization Assessment		0.116 (0.013)	0.081 (0.014)
Female X Polarization Assessment		-0.093 (0.018)	-0.078 (0.019)

Notes: N = 32,152 responses with a non-missing educational attainment and a respondent's assessment of the parties' issue positions on at least one policy domain. The dependent variable is partisan identification with Democrats coded as 0, Independents and Leaners coded as 1/2, and Republican coded as 1. 7-point assessments of a respondent's assessment of the party's issue positions are recoded so that they range from 0 ("Most Liberal") to 1 ("Most Conservative"). "Polarization Assessment" is constructed by subtracting the respondent's average placement of the Democratic party's issue positions from the respondent's average placement of the Republican party's issue position. All regressions also include year fixed effects and observations are weighted by their sample weight. Robust standard errors are reported in parentheses.

Consistent with our expectations, column 2 of Table 2 shows that the partisan gender gap is also larger among those who perceive greater polarization. This is illustrated by the significant negative interaction between the female indicator and the respondent's polarization assessment. Moving from the 25th to the 75th percentile in polarization assessments (.000 to .458) almost doubles the partisan gender gap from -.024 (std. err. = .006) to -.045 (std. err. = .005).<sup>31</sup>

The results in column 3 of Table 2 show that the relationship between college education and the partisan gender gap attenuates, but does not entirely go away, once we account for a respondent's polarization assessment. The coefficient on the interaction between being female and being a college graduate drops about 40% from -0.036 to -0.022 when controls for polarization assessments are also included in the regression. The coefficient on the interaction between being female and polarization assessments also declines slightly from -0.093 to -0.078 when educational

<sup>31</sup>One concern is that relationship may be driven by similar time trends in both the partisan gender gap and polarization assessments. To account for this concern, Table A.19 reports the results of similar regressions that also include female by year fixed effects. The fact that we estimate a nearly identical interaction between female and polarization assessments when these additional fixed effects are included suggests that the relationship between polarization assessments and the partisan gender gap is not simply an artifact of common time trends in these two series.

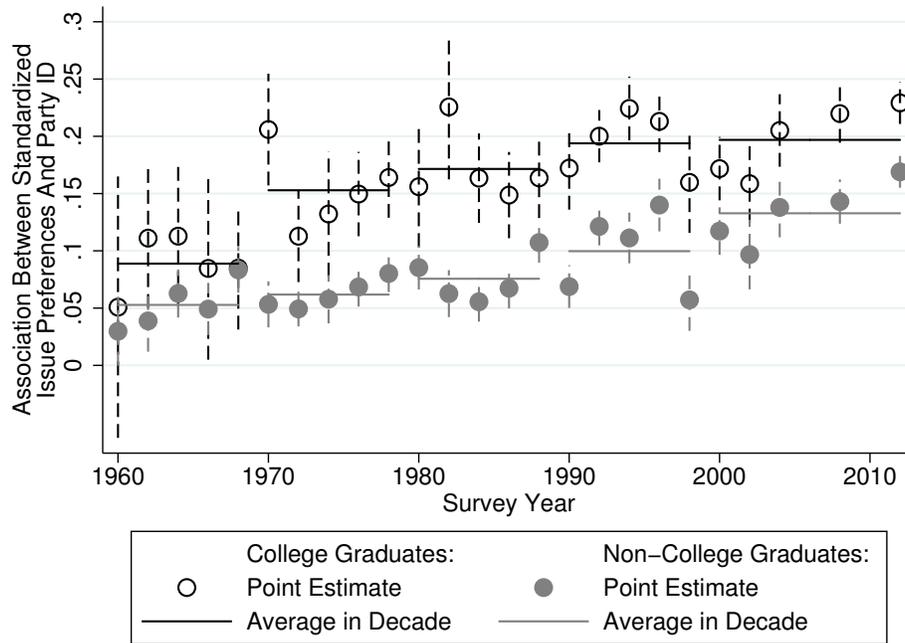
attainment is included in the regression, although the interaction between female and polarization assessments is still significant. Put together these results are mostly, but not entirely, consistent with H2b. The partisan gender gap significantly increases as assessments of polarization increase and these polarization assessments are able to explain a significant portion, although not all, of the relationship between education and partisan gender gap.

### **5.3 Preferences and Partisanship Come Into Alignment When the Gender Gap Grows**

Hypothesis #3 (H3) predicts that, as people become more aware of partisan polarization, both men and women should hold issue positions that better match their partisanship. This occurs because some people adopt the issue positions of their existing parties, while other people sort into parties that better match their preferences (Levendusky, 2009, ch. 6). The second of these two processes led the partisan gender gap to emerge. We contend that it emerged earlier, and remained larger, among the college educated because college graduates were aware of increasing party polarization earlier than non-college graduates, and thus more likely to sort into parties that better matched their preferences. If this is correct, we expect to observe an increasing congruence between the issue positions and party affiliations of college graduates, relative to non-college graduates, over the time period that the partisan gender gap first emerges.

We test H3 in ANES data by examining how the association between party identification and issue preferences varies over time for college graduates and non-college graduates. The dependent variable in these analyses is a respondent's partisan identification, with leaners coded as Independents. The independent variable is an index of the conservatism of a respondent's issue positions. This index of issue positions is constructed with all available issue positions in the ANES cumulative file for each year, where each issue position is rescaled so that 0 is the most liberal response and 1 is the most conservative response. All of the available issue positions are averaged, and then standardized, so that a coefficient represents the change in the probability of identifying as a Republican from a one-standard-deviation increase in conservatism. We run separate regression

Figure 6: Associations between Issue Positions and Partisan Identification Over Time and by Educational Attainment (ANES, 1960 - 2012)



Notes: This figure shows the differences in the estimated regression coefficient when a respondent’s partisan identification is regressed on an index of a respondent’s issue positions in different years. Partisan identification is coded as Democrats as 0, Independents and Leaners as 1/2, and Republicans as 1. The index of issue positions is constructed with all available issue positions in the ANES cumulative file for that given year, where each issue position is rescaled so that 0 is the most liberal response and 1 is the most conservative response. All of the available issue positions are averaged, and then standardized. Thus, a coefficient corresponds to the change in the probability of being Republican from a one-standard-deviation increase in the conservatism of a respondent’s issue positions.

models for college graduates and non-college graduates for every ANES survey between 1960 and 2012.<sup>32</sup>

Consistent with our theory, Figure 6 shows a substantial increase in the association between issue preferences and partisanship among college graduates over the exact same time interval that the partisan gender gap first emerged. The white circles in Figure 6 show the association between issue preferences and college graduates’ partisanship in a given year, while the grey circles show the same association for non-college graduates. In 1960, for example, a one-standard-deviation

<sup>32</sup>One limitation of this analysis is that different issue preference questions are asked in different years. Thus, some of the variation in regression coefficients over surveys likely reflects differences in the specific issue questions asked across years.

increase in the conservatism of issue preferences corresponded with a five and three percentage point increase in Republican identification among college and non-college graduates, respectively. The black and gray horizontal lines present the average of this association for college and non-college graduates, respectively, by decade. In the 1960s, for example, the average association for college graduates was nine percentage points, as compared to five percentage points for non-college graduates. This association increased for college graduates in the 1970s and 1980s much more so than for non-college graduates. This finding is consistent with H3 and our argument that college graduates were sorting more than non-college graduates over this time period.

We conclude this section by investigating on which issues people were sorting. We disaggregate the general issue position index that we used in the previous analysis into three distinct indices: social welfare preferences, gender role preferences, and all other preferences.<sup>33</sup> This disaggregation is motivated, in part, by Kaufmann and Petrocik's (1999) claim that the increased salience of social welfare attitudes, particularly among men, was one of the primary reasons why the partisan gender gap emerged, as well as Inglehart and Norris's (2003) hypothesis that beliefs about gender equality became more important in structuring female's partisan identification over time. To test whether men and women put different weights on these dimensions, we regress party identification on these three indices plus these indices interacted with a female dummy variable.

Table 3 shows that the increased association between party and issue positions is largely driven by changes in the relationship between social welfare preferences and partisan identification over time. Comparing columns 1 and 4 reveals that the association between social welfare preferences and party identification more than doubled between the 1970s and 2000s. While the association between other preferences and party identification also increased, it did so at a much slower rate. This finding is broadly consistent with Kaufmann and Petrocik's (1999) conjecture that the increased salience of social welfare preferences in partisan identification made existing gender differences in social welfare preferences more consequential. Contrary to Kaufmann and Petrocik's argument, but consistent with ours, there is little difference in how issue preferences relate to men's and

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<sup>33</sup>Because the gender role issue position question only was asked in 1972-1984, 1988-2000, 2004, and 2008, our analysis is constrained to these years.

Table 3: Associations between Issue Positions and Partisan Identification Over Time by Issue Position Type and Gender (ANES, 1972 - 2008)

	(1)	(2)	(3)	(4)
Years	1972-1978	1980-1984, 1988	1990-1998	2000, 2004, 2008
N	7341	6377	7713	3212
Female	-0.010 (0.009)	-0.030 (0.010)	-0.033 (0.009)	-0.024 (0.015)
Social Welfare Preferences	0.058 (0.008)	0.064 (0.008)	0.111 (0.007)	0.126 (0.012)
Gender Roles Preference	-0.009 (0.007)	-0.006 (0.008)	0.012 (0.007)	-0.015 (0.011)
All Other Preferences	0.044 (0.008)	0.053 (0.008)	0.048 (0.007)	0.068 (0.012)
Female X Social Welfare Preferences	0.006 (0.010)	0.004 (0.011)	-0.006 (0.010)	-0.009 (0.018)
Female X Gender Roles Preference	0.010 (0.010)	0.027 (0.010)	0.001 (0.009)	0.036 (0.015)
Female X All Other Preferences	-0.008 (0.011)	0.003 (0.011)	-0.005 (0.010)	0.009 (0.017)

Notes: The dependent variable is partisan identification with Democrats coded as 0, Independents and Leaners coded as 1/2, and Republican coded as 1. All available issue positions in the ANES cumulative file are rescaled so that 0 is the most liberal response and 1 is the most conservative response. The average social welfare preference is the mean of the rescaled responses to VCF0805, VCF0806, VCF0808, VCF0809, VCF0830, VCF0839, VCF0886, VCF0887, VCF0889, VCF0890, VCF0893, and VCF0894. The gender roles preference comes from the rescaled response to VCF0834. The average preference on all other positions is the mean of the rescaled responses to all other issue position questions. All three of these measures are standardized. Thus, a coefficient corresponds to the change in the probability that a respondent identifies as a Republican from a one-standard-deviation increase in the conservatism of their issue position on a given domain. All regressions also include year fixed effects and observations are weighted by their sample weight. Robust standard errors are reported in parentheses.

women’s partisan identification. While we need to be cautious drawing conclusions about causal direction from this type of analysis because of possible endogeneity, the similar associations between preferences and party identification among men and women is at least inconsistent with claims that men place greater weight on social welfare preferences.<sup>34</sup>

Table 3 also provides some support for the claim that beliefs about gender roles became more consequential for women’s partisanship over time. For males, there is no decade in which we estimate a statistically significant relationship between gender roles and partisanship. But in the 1980s, 1990s, and 2000s, females with more conservative views about gender roles are significantly more likely to be Republican. Moreover, in the 1980s and 2000s, but not in the 1990s, the association between gender roles is the same for men and women at conventional levels of significance. This finding provides some support for theories, like the Developmental Theory of Gender Realignment, that posit that the salience of gender identity in partisanship has increased over time. However, it is also important to note that the magnitude of this relationship is relatively small. A one standard-deviation increase in conservatism about gender roles associates with about a two percentage point increase in identifying as a Republican. Given that we show in the next section that men’s and women’s preferences over gender roles are becoming more alike over this time period, this means that gender role preferences can explain only a small portion of the development of the partisan gender gap.

## 5.4 Ruling Out Alternative Explanations

### 5.4.1 Generational Replacement

The previous section shows that issue preferences and partisan identification became better aligned over time, particularly among the college educated. We claim that having a college education serves as a proxy for greater awareness of elite polarization. But what if college education is serving as a

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<sup>34</sup>Using Sapiro and Conover’s (1997) terminology, our account of the emergence of the partisan gender gap is a “positional explanation,” meaning that we claim it is the different positions men and women take on issues that is the driving factor, not a “structural explanation,” in which men and women weigh different considerations when making political choices. Sapiro and Conover find some evidence for both explanations driving the gender gap in 1992 U.S. presidential voting. Our claim is that the positional explanation better accounts for long-term partisan change.

proxy for other things instead? The nature of college (and its student population) changed during the twentieth century. About 25 percent of people born in the mid-twentieth century graduated from college, compared to only about 5 percent of people born near the start of it (Goldin, Katz and Kuziemko, 2006). Educational opportunities especially grew for women, as many colleges became coeducational and/or made the curriculum they provided to women more equal to that of men. The Development Theory of Gender Realignment argues that the partisan gender gap results from the socialization women received after the transformation of sex roles, much of which may have happened at college (Inglehart and Norris, 2000). Over time, women’s equality became a larger part of the curriculum and socialization that students received at many colleges. Is change in the types of people attending college, and in the socialization they received there, the reason why a partisan gender gap emerged first, and is consistently larger, among the college educated?

Cohort analysis is one way to differentiate the predictions of our theory from these alternative explanations.<sup>35</sup> If awareness of polarization is the reason why the partisan gender gap emerged first among the college educated, we should expect a partisan gender gap to emerge across all cohorts of college graduates. In contrast, if it is the changing nature of college driving this effect, the partisan gender gap should emerge primarily within the younger cohorts of college graduates. An advantage of our Gallup dataset is that we have enough data to track the birth cohorts over time to test these rival predictions.

Figure 7 presents the partisan gender gap over time separately for those born in each decade from the 1880s to the 1960s for college graduates (top panel) and those with less education (bottom panel).<sup>36</sup> If the growth of the gap among the college educated was driven by generational replacement, in the top panel the lines would be flat and each younger generation’s line would be below the previous generation’s. But that is not what we see. Instead, the lines representing all generations slope down together over time. Except for those born in the 1890s, the partisan gender gap grew on a similar gradual trend among college graduates born in each decade. Even though those who attended college throughout the twentieth century were selected in very different ways,

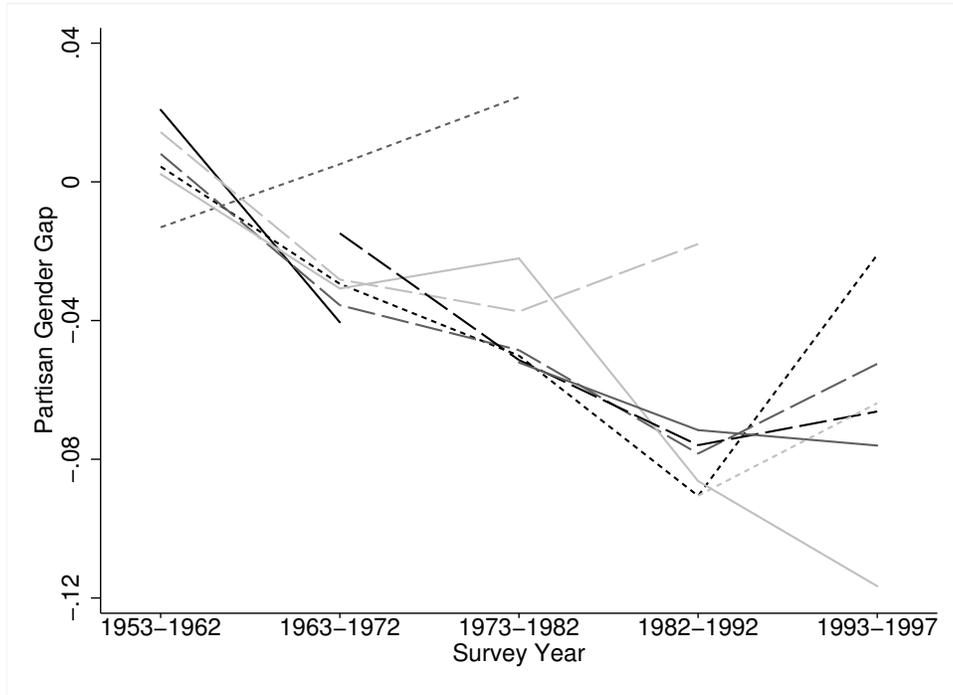
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<sup>35</sup>See Clark (2017) for a cohort analysis of gender differences in policy attitudes.

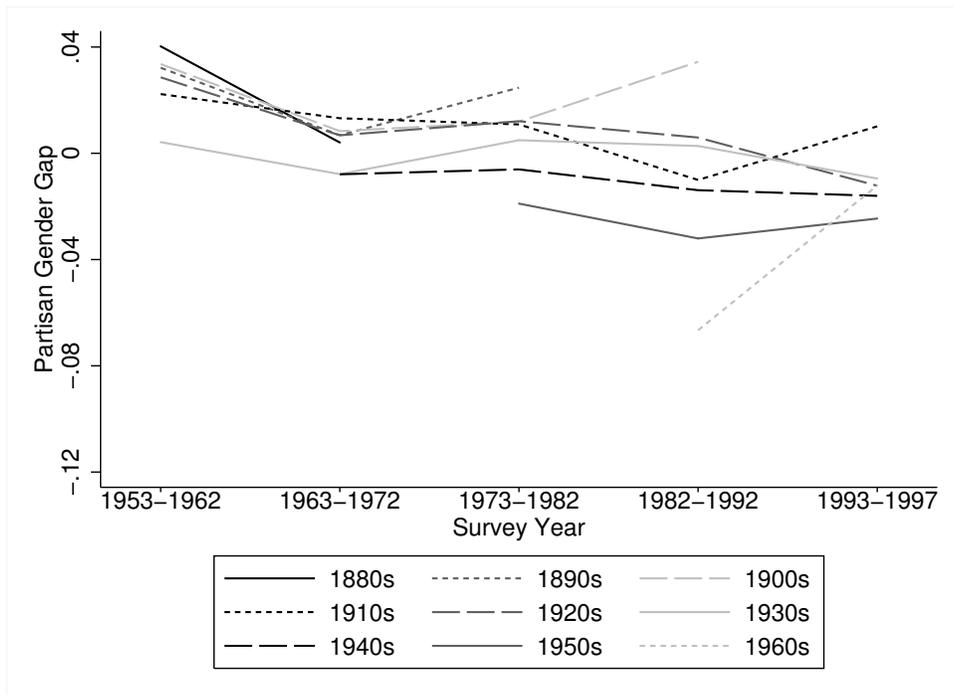
<sup>36</sup>For those who would like to examine the standard errors on these estimates, Table A.20 in the Appendix Section 7.1 provides the data from Figure 7 in table form.

Figure 7: Partisan Gender Gap by Birth Cohort Across Time (Gallup)

(a) College Graduates



(b) Non-College Graduates



and had very different college experiences, the partisan gender gap was similarly sized and evolved in a similar manner among all these cohorts.

Among the non-college graduates, there is some evidence of generational replacement. Between 1983-1992, Figure 7 reveals there was a small modern partisan gender gap among people born in the 1940s or after, which was not present among people born earlier. However, even this doesn't hold up if you extend the time series forward. Table A.13 and Figure A.6 in Appendix Section 7.7, which extend this analysis through 2012 using phone survey data, shows that the differences between these cohorts largely disappeared over time. By the 2000s, men without college degrees were roughly four points more Republican than women without college degrees, even in cohorts where no partisan gender gap was present in the 1980s and early 1990s.

The emergence of the partisan gender gap does not appear to be driven primarily by generational replacement. Rather, it emerged earlier, and is consistently larger, among the college educated of *all* cohorts. We argue that this is because the college educated were more likely to be aware of the growing elite ideological polarization of the parties, understand its significance, and sort into a party that better matched their policy preferences.

#### **5.4.2 Growing Gender Differences in Policy Preferences**

Another alternative explanation for the emergence of the partisan gender gap is a growing divergence in men's and women's policy preferences, not more sorting based on preexisting and steady levels of gender policy differences. As discussed in Section 2, several existing explanations of the U.S. gender gap make growing gender differences in policy preferences a key part of their theories. To examine this possibility, Figure A.5 in Appendix Section 7.6 presents a plot of the gender gap on all issue position questions contained in the cumulative ANES dataset over time. It reveals, consistent with the previous literature reviewed in Section 2, that there are substantial policy preference differences between men and women that pre-date the emergence of the modern partisan gender gap. If anything, the aggregate issue preference gender gap was smaller in the 1970s and 1980s, when the partisan gender gap first emerged, than in the 1960s, 1990s, or 2000s. While one should interpret these results cautiously because the same issue position questions are

not be asked over time, the evidence seems most consistent with our story—that partisan sorting based on preexisting gender policy differences that persisted through this period led to the gradual emergence of the partisan gender gap. Appendix Section 7.6 also looks at specific issues areas and perceptions of economic performance, as well issue preferences separately among those with and without college educations. In none of these cases, do we find evidence that the growth in the partisan gender gap corresponds with substantial growth in preference (or economic perception) gender differences.

## 6 Conclusion

Since the dawn of modern polling in the U.S., men have consistently held more conservative views than women on social welfare, foreign policy, and racial issues. These preference gaps have remained relatively consistent over time. Starting in the 1960s, Republican and Democratic Party elites increasingly diverged ideologically. As people gradually perceived this divergence, they slowly adjusted their party identification to better match their policy preferences. The pre-existing and persistent gender preference differences caused more men to become Republican and more women to become Democrats. This paper tests three hypotheses implied by this argument and finds the evidence to be largely consistent with all three.

An important qualification to our explanation is that ideological sorting does not explain anywhere close to all the variation in macropartisanship over time. There are powerful overall trends that influence macropartisanship among all genders and education levels (Erikson, MacKuen and Stimson, 2002). However, those overall trends don't preclude a secular gender divergence or our explanation of it. Also, while our theory is that ideological party sorting explains the gradual long term growth of the partisan gender gap from the 1960s to the 1990s, it does not explain all historical variation in it. For instance, our theory does not explain how there came to be a small reverse partisan gender gap in the 1950s. In general, while ideological sorting caused most of the secular growth in the partisan gender gap over the late 20th century, we don't claim that it is the only thing influencing the size of this gap in these or other decades.

How does our evidence fit with the literature’s other explanations of the partisan gender gap’s emergence? In Section 2, we identified three major existing groups of explanations for the gender gap’s rise: female economic vulnerability, female labor force participation, and the increasing influence of feminism. Here, we return to each of these.

Much of the evidence that we uncover is inconsistent with the economic vulnerability explanation. The partisan gender gap first emerged because college-educated women became more Democratic in the 1970s. These women, on average, have the most human capital and are at less financial risk than other women from macroeconomic downturns, divorce, or other hardships. Over this time period, we do not observe growing overall gender differences in social welfare preferences, nor do we observe larger gender differences in social welfare attitudes among college graduates than among non-college graduates. Gender differences in assessments of personal economic well-being also remained relatively constant.

Our analysis does not speak as directly to the question of whether increased female labor market participation and economic independence caused the partisan gender gap to grow. This is partially an issue of data availability. Gallup did not consistently ask about labor force participation or marital status until the late 1970s, after the modern partisan gender gap had emerged. However, controlling for labor force participation and marital status, once these variables are available, has little effect on the estimated difference in the partisan gender gap between the college educated and non-college educated. Thus, as best as we can tell with these data, it does not appear that differential labor market participation or marriage rates between the college educated and the non-college educated are what caused the partisan gender gap to emerge earlier, and remain larger, among the college educated.

Our evidence is only partially consistent with explanations based on the feminist socialization of younger generations of women, such as the Development Theory of Gender Realignment. Among college graduates, who are the leading drivers of the partisan gender gap, the gap does not emerge through generational replacement. Among non-college graduates, a small partisan gender gap emerged among younger, but not older, cohorts in the 1980s and 1990s, which could be consistent with Inglehart and Norris’ (2000) argument that changes in sex roles shape the political preferences

of women socialized later. But the magnitude of this partisan gender gap is smaller than the gap that was present in just about all cohorts of college graduates over the same time periods. And by the 2000s, the generational differences largely disappeared among non-college graduates. These findings bring into question the importance of changes in socialization over time as a cause of the partisan gender gap's emergence, at least in the United States.

There is also only limited evidence that liberal views on gender roles or post-materialist social issues caused the partisan gender gap. Among college graduates, men's issue preferences on gender roles have been more conservative than women's since these issue positions were first surveyed on the ANES in the 1970s. But the association between gender role preferences and party identification is small. And a gender gap in opinions about gender roles doesn't appear in the broader population until the 2000s, well after the modern partisan gender gap emerged.<sup>37</sup>

Moving beyond these three existing groups of explanations, the contention by some activists in the early 1980s that the right turn in the policy positions of the national Republican Party created the partisan gender gap has substantial truth to it. However, a simple version of this story risks neglecting the fact that this right turn was a continuation of the gradual ideological polarization of the two parties that began in the 1960s and 1970s. As this polarization continued and grew in size, more people noticed it and the partisan gender gap became larger and somewhat more widespread. Also, it is important to remember that it was elite polarization and mass-level sorting based largely on issues less directly connected with gender that were the main drivers of this phenomenon, as those are the issues with large gender preference differences.

Our analysis helps to bridge the divide between the literature on the partisan gender gap and the most prominent general theories of public opinion and party identification. One of the most prominent scholarly claims about American public opinion is that people with different levels of political awareness comprehend politics and respond to new political information differently (e.g., Converse, 1964; Zaller, 1992). Another major claim is that, because party identification is a social identity, it changes gradually, even in the face of substantial new information (e.g., Campbell et al., 1980; Green, Palmquist and Schickler, 2002). The fact that party identification gradually

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<sup>37</sup>See Table A.21 in Appendix Section 7.4.

responded to party polarization, and did so differently depending on education, is consistent with these two seminal arguments. New theories of mass political behavior are not required to explain the development of the modern partisan gender gap. Rather, its formation is an understandable consequence of men’s and women’s divergent policy preferences, the polarization of the party system, and differences in how those with different levels of political awareness learned about and responded to this changing political landscape.

Our results also illustrate how mass-level ideological sorting can have unintended consequences. Whenever there are preexisting policy preference differences across demographic groups, ideological sorting can enlarge demographic differences between the parties. Relatedly, when national political parties demographically diverge, it could be caused by any major political issue, not necessarily by explicitly group-based political appeals.

Finally, our study, and the survey data we have assembled, open the door for more research on this topic. For instance, our examination of the interaction between gender, party identification, and education with other demographic variables has so far been limited. In 1 in Section 5.2, we show that the relationship between education and the partisan gender gap persists even when controlling for a series of demographics. And as noted in Section 5.2, Tables A.14 through A.17 in Section 7.8 in the Appendix show how the gender gap grows over the years within a series of demographic categories in addition to education, such as age, race, religion marital status, union membership, household income, employment status, and place of residence. While these tables corroborate H2a—that the biggest and most consistent demographic difference is between the more and less educated—they also suggest avenues for future work. It would be great if future research further examined variation in the relationship between party identification, gender and education over time within certain other demographic categories, like African Americans, Southerners, the unmarried, etc. While our work has examined the interplay of a series of variables—party identification, gender, education, issue preferences, perceptions of party polarization, time, age—we look forward to future work that takes a closer examination of how the relationships we find depend on other attributes.

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