

A Quantitative Theory of the Credit Score

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Intro



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- In a world where agents have persistent unobservable characteristics, credit scores can help allocate credit among risky borrowers.
- We don't assume exogenous exclusion or stigma following default, but instead people's credit score falls and subsequently face worse borrowing terms.
- Our theory matches the data from unsecured consumer credit markets and the age profile of credit scores.



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 - Our baseline economy is also preferred by the median newborn to a hidden info economy where credit market behavior is not tracked.



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 - Deliver tractability.



- ① Model
- ② Equilibrium
- ③ Taking the Model to Data
- ④ Model Properties
- ⑤ The Role of Persistent Hidden Information for Credit Access

Model



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- Lenders try to infer a HH's type from their credit market actions, updating their assessments using Bayes' law.
- Lenders use those assessments of creditworthiness to price their loans.



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 - HH cannot save that period ($a' = 0$)
 - and pays filing cost κ ($c = e + z - \kappa$) \rightarrow “static” punishment



- risk neutral, perfectly competitive (free entry)

Inference problem: can't observe β, z or $\epsilon^{(d, a')}$ to price loans:

Solution: assign a “type score”, subjective prior $s = \Pr(\beta = \beta_H)$

Pricing: offer discount loans at prices $q^{(0, a')}(\omega)$, where

$$q^{(0, a')}(\omega) = \begin{cases} \frac{\rho p^{(0, a')}(\omega)}{1+r} & \text{if } a' < 0 \\ \frac{\rho}{1+r} & \text{if } a' \geq 0, \end{cases}$$

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Solution: assign a "type score", subjective prior $s = \Pr(\beta = \beta_H)$

- update via Bayes rule using observables (d, a') and $\omega = (e, a, s)$ to revise posterior score $\psi^{(d, a')}(\omega)$

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- ② HH receive transitory earnings and preference shocks
 - $z \sim H(z)$
 - $\epsilon = \{\epsilon^{(d, a')}\} \sim F(\epsilon; \alpha, \lambda)$
- ③ Given price schedule $q = \{q^{(0, a')}(\omega)\}$, agents choose (d, a')
- ④ Intermediaries revise type scores from $s \rightarrow \psi^{(d, a')}(\omega)$ via Bayes rule.
- ⑤ Next period states are drawn:
 - Individuals who survive with prob ρ draw $\beta' \sim Q^\beta(\beta'|\beta)$, $e' \sim Q^e(e'|e)$, and $s' \sim Q^s(s'|\psi)$.
 - Newborns draw $\beta' \sim G_\beta$, $e' = \underline{e}$, $a' = 0$, and s' consistent with G_β .

Estimation, Demographics, Both



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Equilibrium



Definition

A stationary **recursive competitive equilibrium (RCE)** comprises:

- pricing function $q^*(\omega)$ (vector-valued)
- type scoring function, $\psi^*(\omega)$ (vector-valued)
- choice probability function, $\sigma^*(\beta, z, \omega)$ (vector-valued)
- steady state distribution, $\mu^*(\beta, z, \omega)$

such that

- $\sigma^*(\beta, z, \omega)$ consistent with HH dynamic optimization ▶ HH
- $\psi^*(\omega)$ satisfies Bayes' Rule ▶ score
- $q^*(\omega)$ implies lenders break even, with repayment probabilities $p^{(0,a')}(\omega)$ implied by σ^* ▶ price
- $\mu^*(\beta, z, \omega)$ is a fixed point of $\mu' = T^* \mu$ ▶ Distn



Theorem

Existence: *There exists a RCE.*



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- We show how a RCE with *type scores* s is equivalent to the stationary eqm (RCECS) of an economy where prices depend on *credit scores*.
- Credit scores just the prob. of repayment $m \equiv p^{(0, \bar{a})}(e, a, s)$ on loan \bar{a} .

Theorem

Equivalence: *If there exists a RCE and there is a one-to-one mapping between type scores s and credit scores m , then there exists a RCECS.*

Mapping the Model to Data



- A credit score is an *ordinal* measure of creditworthiness (i.e. What does 760 really mean?).
- To take the theory to data, we associate with $p^{(0, \bar{a})}(\omega) \in [0, 1]$ a number that gives $p^{(0, \bar{a})}(\omega)$'s position (ranking) in the overall distn. (i.e. the fraction of people with lower credit scores)
- We check that rankings preserve order (i.e. do not depend on the specification of \bar{a} , which we take to be 3.5% of median earnings).

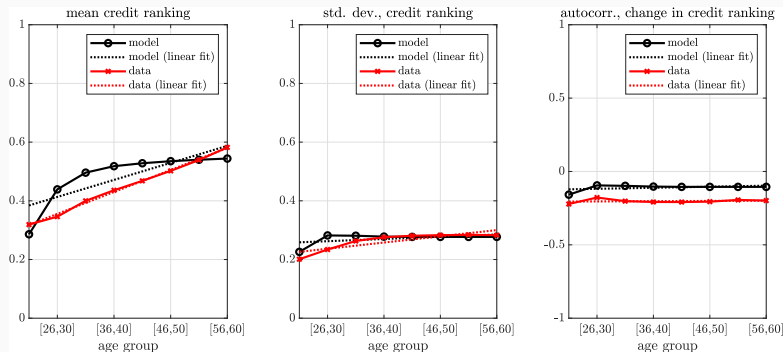


Parameter		Value	Notes (annual, 2 types)
Demographics and preferences			
survival probability	ρ	0.975	avg. 20-60 age profile
risk aversion	γ	3	CRRA preferences
persistent earnings at birth	\underline{e}	0.57	lowest e level
Technology			
risk-free rate (%)	r	1.0%	real annual T-Bill
filing cost	κ	0.02	2% median earnings
debt level for computing credit score	\bar{a}	-0.035	3.5% median earnings
Earnings processes			
persistence of $\log(e)$	ρ_e	0.914	Floden and Linde (2001)
std. dev. of innovation to $\log(e)$	ν_e	0.043	Floden and Linde (2001)
std. dev. of $\log(z)$	ν_z	0.042	Floden and Linde (2001)



Parameter		Value	Moment (%)	Data	Model
Evolution of types			Aggregate credit market statistics		
high β	β_H	0.915	default rate	0.99	0.98
low β	β_L	0.886	avg. interest rate	12.87	13.95
high→low prob.	Q_{HL}^β	0.011	IR dispersion	6.56	7.24
low→high prob.	Q_{LH}^β	0.013	fraction in debt	10.43	10.50
frac β_H newborns	G_{β_H}	0.280	debt to income	0.35	0.25
Extreme value parameters			Credit score life cycle moments		
EV scale (1e-3)	α	3.387	intercept, mean	0.281	0.355
EV correlation	λ	0.991	slope, mean	0.037	0.029
			intercept, std. dev.	0.216	0.255
			slope, std. dev.	0.011	0.004
			avg autocorr of change	-0.202	-0.109

► Identification



- Mean and standard deviation of rankings rise over age
- Autocorrelation of changes fall (i.e. mean reversion of rankings).



About hidden type:

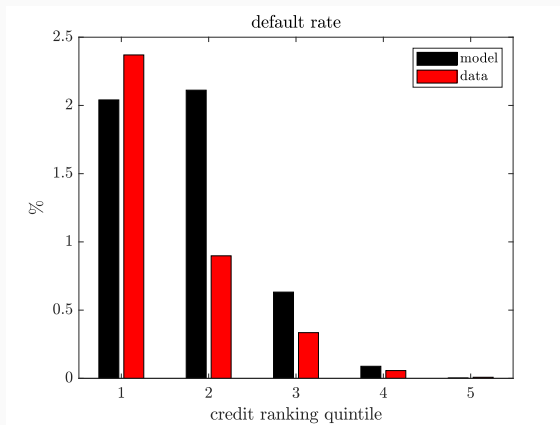
- Data consistent with small type differences $\beta_H = 0.915 > 0.886 = \beta_L$.
- Types are persistent $> 98\%$ and more likely to switch to high.
- Fraction of newborn high types is 0.28, lower than stationary fraction 0.44 and long run 0.56 (i.e. scores grow over time).

About EV shocks:

- Low variance of shocks ($\alpha = 0.003$) and high correlation ($\gamma = 0.99$) implies decisions are concentrated at the modal choice (e.g. share of modal defaulters is 87%).

Insight: Relation between $\beta_H - \beta_L$ and α important for inference problem.

- If $\beta_H - \beta_L$ large, then easy to separate types and would need large α to cloud inference.



- Higher credit rankings have lower default rates.
- Many low earners in model Q2 leading to higher default rates than the data.

Model Properties



- 1 As in the data, earnings, wealth, scores, cross-sectional consumption variance rise with age. ▶ Age
- 2 Type L default and borrow more, save less (i.e. riskier). ▶ DBS
- 3 Default falls with higher e , higher type β , lower debt. ▶ Default
- 4 Dynamic Scoring: scores fall with default and more debt, rise if pay off debt, and are mean reverting. ▶ DynScore
- 5 Adverse Selection: the price of not using dynamic scoring yields a riskier pool of borrowers. ▶ AS
- 6 Higher scores lower interest rates for given debt and more debt makes rates more sensitive to one's score. ▶ Rates
- 7 Conditional on default, interest rates rise. ▶ drates
- 8 Signalling: Type L mimicking type H is costly & grows. ▶ Mimic

Role of Information for Credit Access



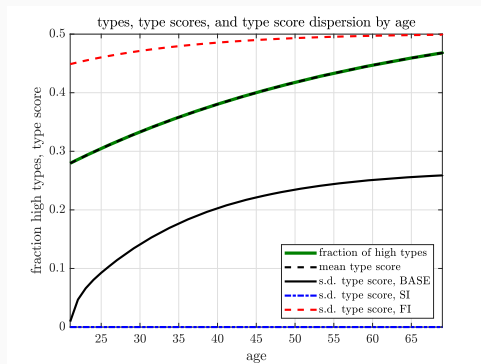
Compare with two alternative info structures:

① Full information (FI)

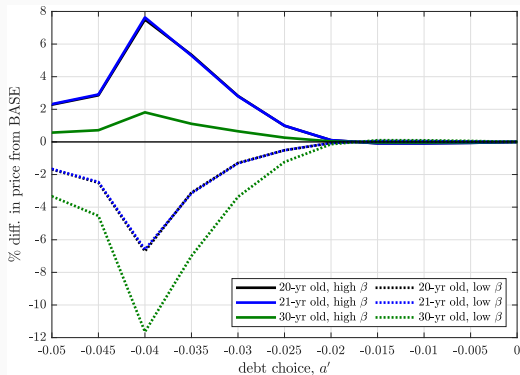
- Type is observable \implies no inference problem \implies no type scores.
- History does not matter for pricing as in CCNR (2007).
- Actual type directly affects prices, $q_{FI}^{(0,a')}(\beta, e)$.

② No tracking (NT)

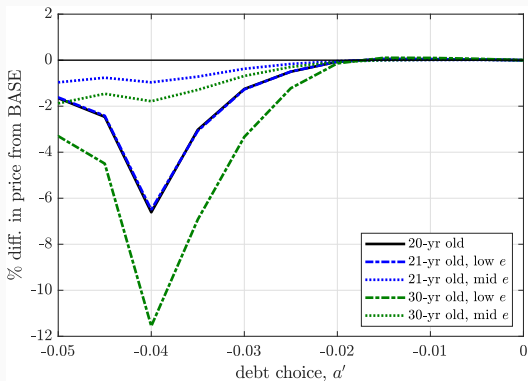
- Lenders cannot use the history of actions to track type so type score only reflects one's age (length of credit history).
- Incentives are weakened for good behavior but pooling provides partial insurance (tradeoff \longrightarrow quantitative question).
- Price function $q_{NT}^{(0,a')}(age, e)$.



- All info structures have the same mean score given by fraction of H types:
 $\bar{s}' = \bar{s} \cdot Q^\beta(H|H) + (1 - \bar{s}) \cdot Q^\beta(H|L)$ over one's life with estimated initial fraction $G_{\beta H} = 0.28$.
- Dispersion in scores weakly rise with age.
- No dispersion in NT (where type histories are pooled so nothing is learned) and highest dispersion in FI (because they are perfectly separated).



- Type H face higher prices (lower rates) in FI than BASE especially when young.
- Riskier type L face lower prices (higher rates) in FI than BASE especially when old.



- Lack of incentives to build a reputation in NT lead to higher interest rates than BASE.
- As time passes, earnings and adverse selection play more of a role; if an individual wants to borrow a lot, they are more likely to be type L who have not accumulated enough wealth (so face higher rates).



model subgroup	% difference relative to BASE					
	No Tracking (NT)			Full Info (FI)		
	high β	low β	all	high β	low β	all
volume of debt	-4.20	-3.16	-3.50	+6.28	-3.14	-0.08
debt to income	-4.02	-2.97	-3.32	+5.77	-2.95	-0.09
default rate	-9.03	-6.85	-7.52	+12.2	-6.84	-1.00

Important differences:

- Type *H* have more debt in FI (since they face lower rates) and less debt in NT than in BASE (since they are pooled with riskier type *L*).
- Type *L* have less debt (which induces less default) in both FI and NT.
- Aggregate volume of debt and default lower in NT (negative consequences consumption smoothing).



Consumption Equivalent gain relative to BASE for newborns (in %)

model	No Tracking (NT) mean	Full Info (FI) mean
high β	-0.41	+1.04
low β	-0.55	-0.47
all	-0.51	-0.05

Important differences between young (who want to borrow against future income) and old (who have accumulated assets):

- FI: type L hurt, type H helped by more separation. Avg newborn hurt.
- NT: incentive effects dominate pooling, newborns especially hurt (1/2 of 1 percent is big in Lucas terms).

Conclusion



- We show how full info models of unsecured debt with risk of default can easily be extended to handle hidden info via EV shocks to preferences and recursive updating of type scores.



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- Since the young tend to borrow against their future income and the old tend to accumulate precautionary balances, less info (i.e. "small data" NT economy) worsens incentives to repay (raising interest rates) which dominates insurance afforded by partial pooling, yielding lower welfare than a "big data" (BASE) economy.



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- The intersection of information and demographics matters!

Appendix

HH PROBLEM

$$V(\beta, z, \omega, \epsilon) = \max_{(d, a') \in \mathcal{F}(z, \omega)} \underbrace{v^{(d, a')}(\beta, \omega)}_{\text{fundamental value}} + \underbrace{\epsilon^{(d, a')}}_{\sim \text{GEV}}$$

where the conditional value function is given by

$$v^{(d, a')}(\beta, z, \omega) = (1 - \beta\rho)u(c^{(d, a')}(z, \omega)) + \beta\rho \sum_{\beta', z', \omega'} Q^\beta(\beta'|\beta)Q^e(e'|e)H(z')Q^s(s'|\psi^{(d, a')})W(\beta', z', \omega'),$$

budget feasibility $(d, a') \in \mathcal{F}(z, \omega) \iff c^{(d, a')}(z, \omega) \geq 0$ given by

$$c^{(d, a')}(z, \omega) = \begin{cases} e + z + a - q^{(0, a')}(\omega) \cdot a' & \text{if } (d, a') = (0, a') \\ e + z - \kappa & \text{if } a < 0 \text{ and } (d, a') = (1, 0) \end{cases},$$

and the expected future value function is given by

$$W(\beta, z, \omega) \equiv \int V(\beta, z, \omega, \epsilon) dF(\epsilon).$$

Extreme value shocks generate simple choice probabilities:

Probability of Default: If $a < 0$,

$$\sigma^{(1,0)}(\beta, z, \omega) = \frac{\exp \left\{ v^{(1,0)}(\beta, z, \omega) / \alpha \right\}}{\exp \left\{ v^{(1,0)}(\beta, z, \omega) / \alpha \right\} + \exp \left\{ \lambda W_{ND}(\beta, z, \omega) / \alpha \right\}}$$

and 0 otherwise, where

$$W_{ND}(\beta, z, \omega) = \alpha \log \left(\sum_{(0, \bar{a}) \in \mathcal{F}(z, \omega)} \exp \left\{ v^{(0, \bar{a})}(\beta, z, \omega) / \lambda \alpha \right\} \right)$$

Conditional on not defaulting, a HH chooses $(0, a') \in \mathcal{F}(z, \omega)$ via via

$$\tilde{\sigma}^{(0, a')}(\beta, z, \omega) = \frac{\exp \left\{ v^{(0, a')}(\beta, z, \omega) / \lambda \alpha \right\}}{\sum_{(0, \bar{a}) \in \mathcal{F}(z, \omega)} \exp \left\{ v^{(0, \bar{a})}(\beta, z, \omega) / \lambda \alpha \right\}},$$

with $\tilde{\sigma}^{(0, a')}(\beta, z, \omega) = 0$ if $(0, a') \notin \mathcal{F}(z, \omega)$. [▶ Back](#)

- Type Scores from an action $(d, a') \in \mathcal{F}(z, \omega)$ are updated via Bayes' Rule via:

$$\begin{aligned}\psi_{\beta'}^{(d, a')}(\omega) &\equiv \Pr(\beta' | (d, a'), \omega) \\ &= \sum_{\beta} Q^{\beta}(\beta' | \beta) \frac{\sum_z \sigma^{(d, a')}(\beta, z, \omega) H(z) s(\beta)}{\sum_{\tilde{\beta}, z} \sigma^{(d, a')}(\tilde{\beta}, z, \omega) H(z) s(\tilde{\beta})}\end{aligned}$$

LOAN PRICING AND REPAYMENT PROBABILITIES

- The prob. of repayment on a loan a' given to an agent in observable state ω is:

$$\begin{aligned} p^{(0,a')}(\omega) &\equiv \Pr(d' \neq 1 | a' < 0, \omega) \\ &= \sum_{\beta', z', e', s'} H(z') \cdot Q^e(e'|e) \cdot Q^s(s'(\beta') | \psi_{\beta'}^{(0,a')}(\omega)) \cdot s'(\beta') \\ &\quad \cdot (1 - \sigma^{(1,0)}(\beta', z', e', a', s')). \end{aligned}$$

- The price function is:

$$q^{(0,a')}(\omega) = \begin{cases} \frac{\rho p^{(0,a')}(\omega)}{1+r} & \text{if } a' < 0 \\ \frac{\rho}{1+r} & \text{if } a' \geq 0, \end{cases}$$

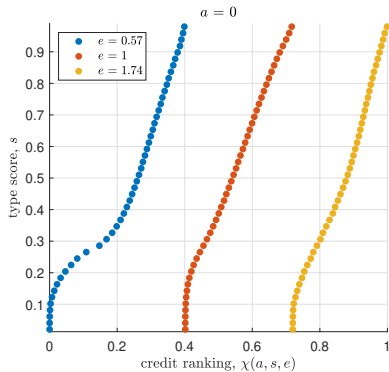
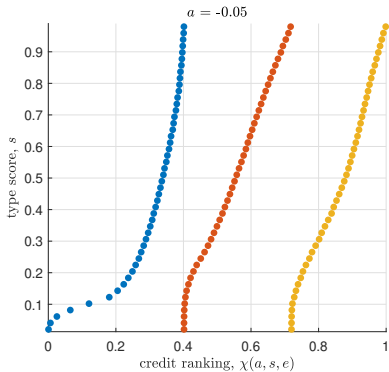
$$\mu'(\beta', z', \omega') = \sum_{\beta, z, \omega} T(\beta', z', \omega' | \beta, z, \omega) \cdot \mu(\beta, z, \omega), \quad (1)$$

where

$$\begin{aligned} T(\beta', z', \omega'; \beta, z, \omega) = & \quad (2) \\ & \rho \cdot Q^\beta(\beta' | \beta) \cdot H(z') \cdot Q^e(e' | e) \cdot \sigma^{(d, a')}(\beta, z, \omega) \cdot Q^s(s'(\beta') | \psi_{\beta'}^{(d, a')}(\omega)) \\ & + (1 - \rho) \cdot G_\beta(\beta') \cdot H(z') \cdot G_e(e') \cdot \mathbf{1}_{\{a'=0\}} \cdot \mathbf{1}_{\{s'=G_\beta\}}. \end{aligned}$$

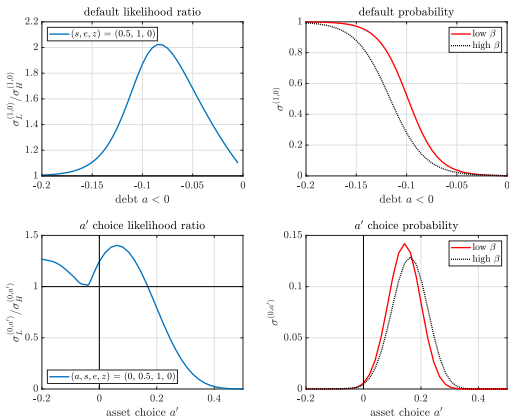
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MAPPING BETWEEN TYPE SCORES AND CREDIT SCORES



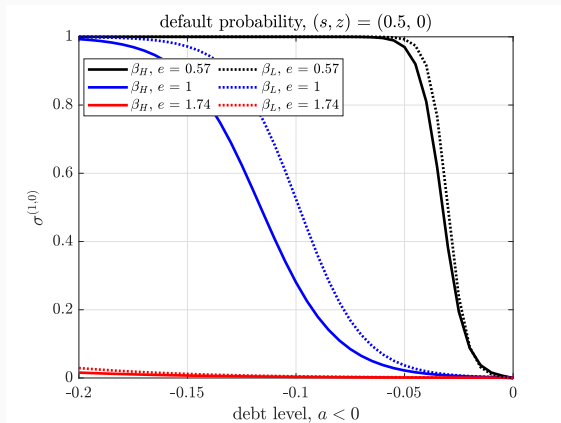
- One-to-one (invertibility satisfied).
- Low earners have low rankings and current assets only important for “subprime” low earners.

DEFAULT AND BORROWING / SAVING DECISIONS



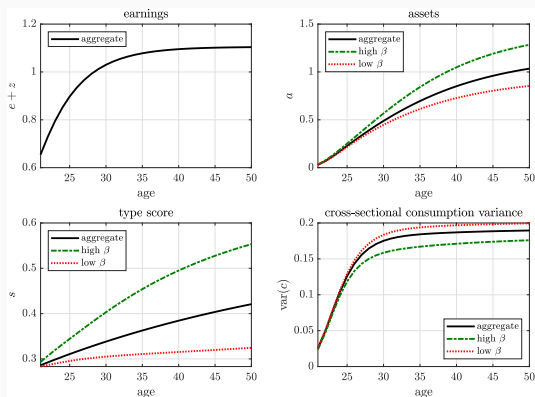
- β_L more likely to default and borrow, less likely to save.
- Non-monotonicity in debt choice occurs at actions w.p. 1^{-8} .

FACTORS IN DEFAULT



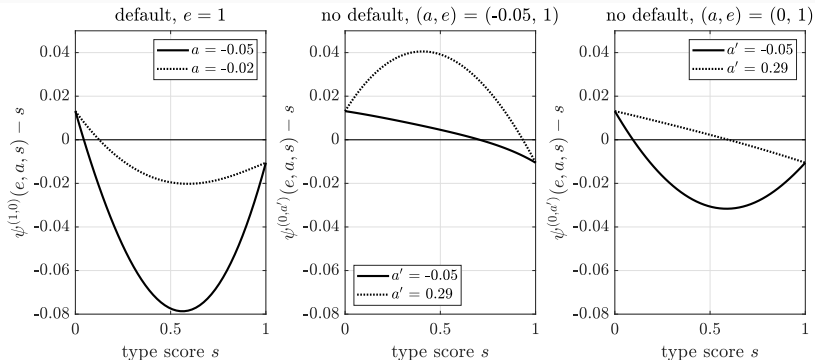
- Default falls with higher e (critical), higher type β , lower debt.

Figure 1: Moments by Age in Baseline Model



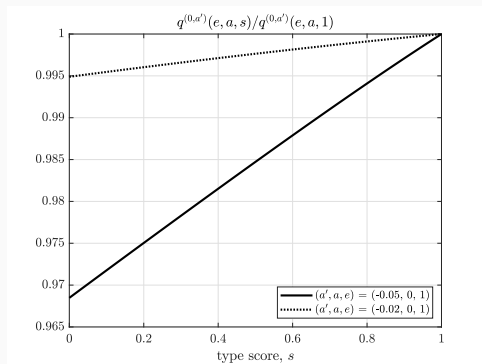
- Earnings induces wealth, scores induce cross-sect. cons. var. (rising c.c.v. age profile as in data - HPV(2010))

DYNAMIC SCORING RESPONSES



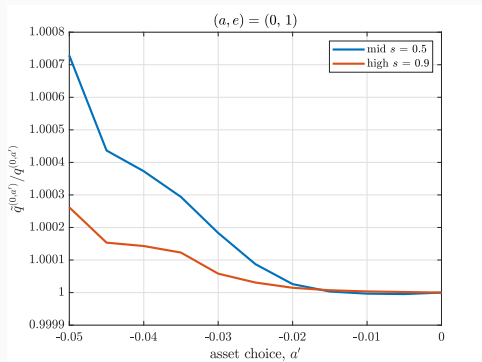
- Real world credit scores: (i) are mean reverting (all) (ii) fall in bankruptcy (lhs); (iii) rise if one pays off one's debt (mid), and (iv) fall as one takes on more debt (rhs).

REPUTATION AND PRICES



- Relative price of unsecured credit across type score: higher score \rightarrow higher relative price (lower interest rate) and more debt \rightarrow more sensitivity to score.

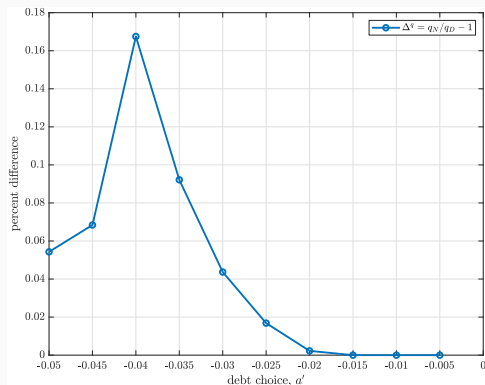
ADVERSE SELECTION EFFECTS



- AS: the bigger the loan, the higher the prob. lenders assign to the borrower being less creditworthy.
- The relative price of not using dynamic scoring yields a riskier (type L) pool of borrowers.

► Back

AVERAGE PRICE IMPACT OF DEFAULT



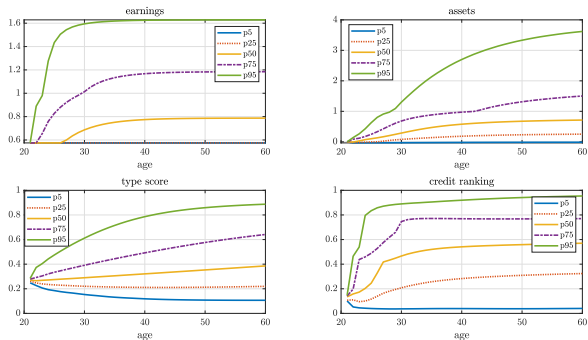
- Given extreme value shocks, there are always two identical agents (i.e. in the same state), one of whom defaults and another who does not.
- Can quantify the impact of default on interest rates by comparing price schedules in the following period conditional on the default choice.
- $q_N - q_D > 0$: rates fall relative to a defaulter.

STATIC SIGNALLING COSTS VERSUS BENEFITS

	% Average Gain in:		
	Consumption (\hat{C})	Wealth (\hat{A})	Reputation ($\hat{\psi}$)
All	-3.78	6.09	5.10
Newborns	-0.84	15.87	3.73

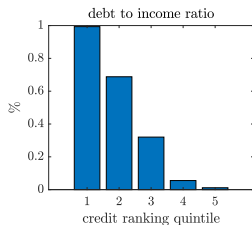
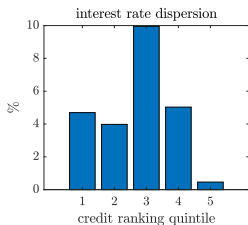
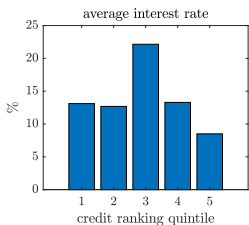
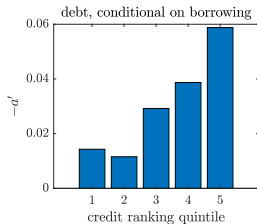
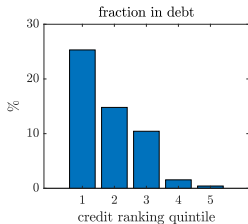
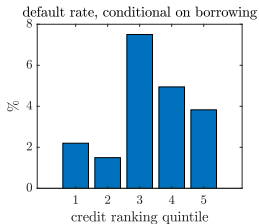
- AS: “bad” types (L) want to mimic “good” types (H) to get better borrowing terms while type H want to separate themselves from L .
- What’s the $\% \Delta$ to consumption (\downarrow) and reputation (\uparrow) if type L follows $\sigma^{(d,a')}(\beta_H, z, \omega)$ instead of $\sigma^{(d,a')}(\beta_L, z, \omega)$?
- Mimicking loss costly for young since there is more pooling.

HOW DO CREDIT RANKINGS VARY BY AGE?

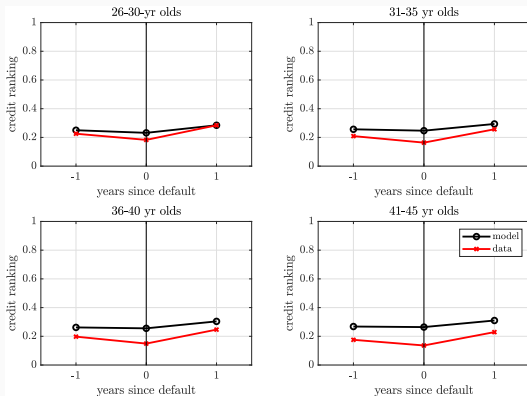


- (tl) earnings grow with age inducing assets growing with age (tr),
- (bl) type score grows with age on average (consistent with $Q_{LH'}^\beta > Q_{HL'}^\beta$) but for some it falls with age (those whose actions reveal they are type L) inducing a similar pattern for rankings (br). ▶ Rank by Outcome
- Takeaway: More pooling when young, more separation when old.

HOW DO CREDIT OUTCOMES VARY BY CREDIT RANKING?



DEFAULT EVENT STUDY (UNTARGETED)



- Bankrupts are in bottom quartile of the credit score distn.
- Decline and subsequent recovery consistent with mean reversion as in Musto (2004).

IDENTIFICATION

While all moments jointly identify the parameters, we find (see sensitivity analysis in Appendix):

- Affine age profile of credit rankings helps identify the type transition process:
 - Constant (and slope) of avg. ranking $\rightarrow G_{\beta_H}$ (and $Q^\beta(H'|L) - Q^\beta(L'|H)$)
 - Constant (and slope) of st. dev. ranking \rightarrow extreme value parameter λ (and Q^β)
 - Neg. autocorrelation of changes (mean reversion) \leftarrow earnings process.
- Aggregate moments help identify:
 - Default rate \rightarrow extreme value parameter α
 - Average interest rate $\rightarrow \beta_H$
 - Interest rate dispersion $\rightarrow \beta_H - \beta_L$

GENERALIZED EXTREME VALUE DISTRIBUTION

GEV cumulative distn. given by:

$$F(\epsilon_t; \alpha, \lambda) = \exp \left\{ - \left[\sum_{a_{t+1} \in \mathcal{A}} \exp \left(- \frac{\epsilon(0, a_{t+1}) - \mu}{\lambda \alpha} \right) \right]^\lambda - \exp \left(- \frac{\epsilon(1, 0) - \mu}{\alpha} \right) \right\}$$

- $\alpha > 0$ governs the variance of the shocks
 - $\uparrow \alpha \implies \uparrow$ variance: decisions more “random”
- $\lambda \in [0, 1]$ determines correlation between shocks associated with non-default actions
 - $\uparrow \lambda \implies \downarrow$ correlation: decisions more “random” *conditional on not defaulting*
 - $\lambda \cdot \alpha$ matters
- $\mu = -\alpha\gamma_E$ set to make the shocks mean zero in expectation